

CMB likelihoods at the frontier

Squeezing parameters out of data

Giacomo Galloni

Challenges at the CMB Frontier · Chino, Japan · July 3, 2026

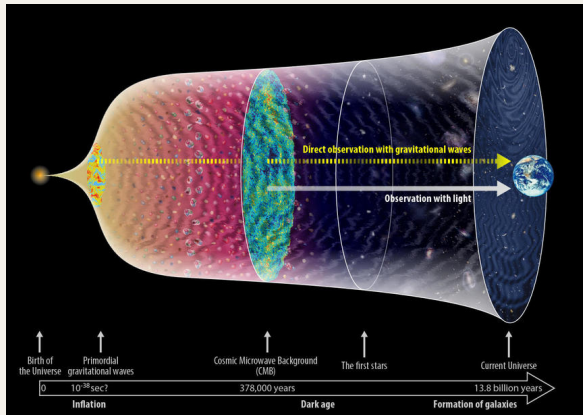
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Inflation — and what it leaves behind

THE QUESTION

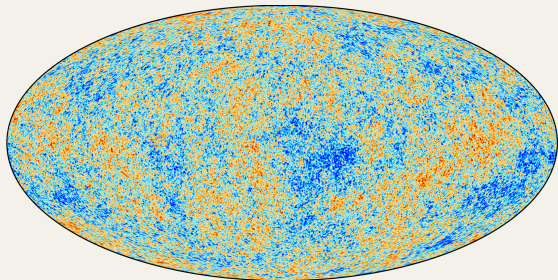
- Inflation is our best framework for the very early universe, yet still **untested in detail**.
- It predicts two kinds of primordial perturbations:
 - **scalar** (density) — the seeds of structure;
 - **tensor** (gravitational waves) — the smoking gun.
- Both leave an imprint on the cosmic microwave background.



The CMB — our window on the early universe

THE QUESTION

- A snapshot of the universe at recombination ($z \approx 1100$).
- Tiny temperature fluctuations $\Delta T/T \sim 10^{-5}$ encode the seeds of all later structure.
- Subsequent reionization and primordial GWs leave their imprints in **polarization**.

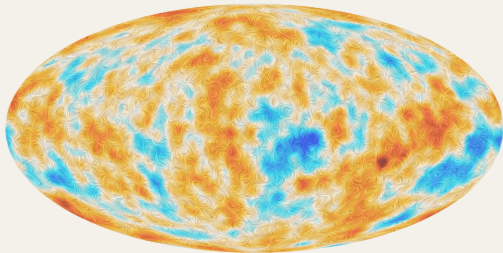


Planck Collaboration — arXiv:1807.06205

CMB polarization — E-modes & B-modes

THE QUESTION

- Linear polarization decomposes into **E-modes** (parity-even) and **B-modes** (parity-odd).
- Scalar perturbations source E-modes; **primordial GWs uniquely source large-scale B-modes**.
- Both the reionization signal and primordial B-modes live at **large angular scales** (low ℓ).



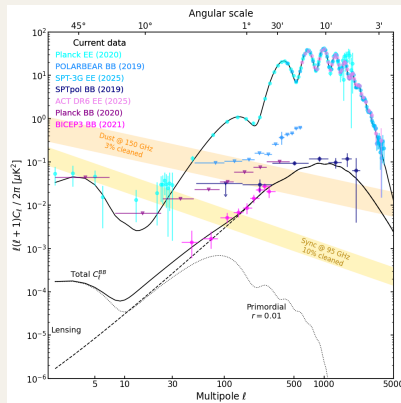
The driving question

WHAT THIS TALK IS REALLY ABOUT

The question

How do we extract information on the parameters of interest from the data?

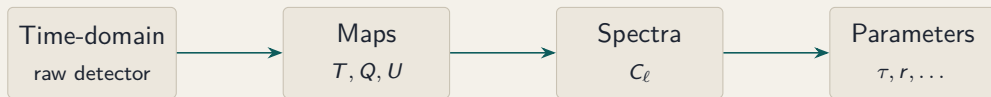
- It hides real complexity: *which* data D , how to **compress** it, how to write a **likelihood**, how to dodge **systematics**.
- This era's challenge — a **complete** CMB description: small scales (not in this talk) *and* large scales.
- Large scales are hard: **no central-limit theorem** for the C_ℓ .
- τ (EE) and r (BB) are our **proxies** here.



E/B-mode measurements to date — our grasp of the large scales is still partial.

The pipeline — sky to parameters via $\mathcal{L}(D \mid \theta)$

THE BRIDGE



Each arrow **compresses** the data — but *how* differs:

- TOD → maps: through a **likelihood** (map-making);
- maps → spectra: **brute** linear compression (spherical-harmonic / QML estimators);
- spectra → parameters: a **likelihood** again — often with crude approximations of the experiment's response.

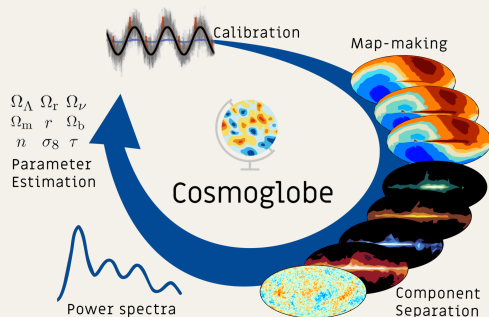
At each step the question is the same: *how much information do we keep, and what do we assume?*

$D =$ time-ordered data

THE RAW STREAM

The detectors deliver a **time-ordered stream** of raw data, tightly coupled to whatever happens in the sky as the instrument scans.

- It is the **most informative** data — every correlation is still in there.
- Working with it directly is the most ambitious likelihood — and computationally enormous; full end-to-end analyses are demonstrated only at limited scale.



Watts et al. — arXiv:2304.05484

Map-making — compressing time into a map

TOD → MAPS · THE STEP WE WERE ASSUMING

Model the calibrated stream as $\mathbf{d} = \mathbf{P}\mathbf{m} + \mathbf{n}$, with $\mathbf{P}$ the pointing matrix (T, Q, U projection) and $\mathbf{N} \equiv \langle \mathbf{n}\mathbf{n}^T \rangle$ the **time-domain** noise covariance. Map-making **compresses the time dimension**: project the years-long stream onto a map.

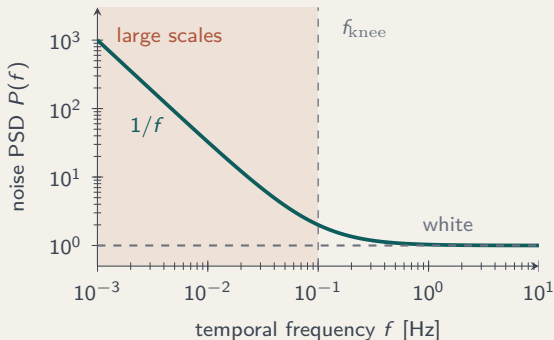
- Simplest, **binning**: average the samples falling in each pixel, $\hat{\mathbf{m}} = (\mathbf{P}^T \mathbf{P})^{-1} \mathbf{P}^T \mathbf{d}$.
- The natural place to remove **some systematics**: cosmic-ray **glitches** (deglitched on the timeline) and **pointing / half-wave-plate wobble** effects. → S. Stellati, Wed.
- The output is a **map** — the data product the rest of the pipeline consumes.

Biquard 2024 — arXiv:2406.15415 · Planck HFI — arXiv:1502.01586.

Why map-making is hard — $1/f$ and the atmosphere

THE NOISE IS NOT WHITE

- Detector noise has a $1/f$ tail: power rises below a *knee* frequency f_{knee} .
- Low temporal frequency $\leftrightarrow$ large angular scales $\leftrightarrow$ low ℓ — the τ /B-mode regime.
- Ground-based: the atmosphere adds a huge, detector-correlated $1/f$ -like signal — the dominant contaminant.
- Naive binning averages as if white $\Rightarrow$ stripes and lost large-scale power.



From binning to the optimal map

GENERALISED LEAST SQUARES

Binning is optimal *only* for white noise. With $1/f$ correlations we must do better:

- **Destriping**: model the $1/f$ part as a few **baseline offsets**, fit and subtract them, then bin — cheap, and recovers most of the optimal answer.
- **Map-making as a likelihood**: maximise the Gaussian TOD likelihood
 $-2 \log \mathcal{L} = (\mathbf{d} - \mathbf{P}\mathbf{m})^\top \mathbf{N}^{-1}(\mathbf{d} - \mathbf{P}\mathbf{m})$. Its **generalised-least-squares** (minimum-variance) maximum and pixel covariance are

$$\hat{\mathbf{m}} = (\mathbf{P}^\top \mathbf{N}^{-1} \mathbf{P})^{-1} \mathbf{P}^\top \mathbf{N}^{-1} \mathbf{d}, \quad \mathbf{C}_N = (\mathbf{P}^\top \mathbf{N}^{-1} \mathbf{P})^{-1}.$$

- It is the **optimal** compression — and **reduces to binning** when the noise is white ($\mathbf{N} \propto \mathbf{I}$).

Challenges of map-making

WHAT MAKES IT HARD

- **Correlated noise, where it hurts.** Instrumental $1/f$ + atmosphere (ground) pile up at low ℓ — the τ /B-mode regime; detector decorrelation worsens the lowest multipoles.
- **Large-scale power loss.** Time-domain filtering / ground removal suppress the biggest scales; unbiased recovery needs a transfer function (filter-and-bin) or full GLS.
- **Data volume.** $\sim 10^4$ – 10^5 detectors $\times$ years, and growing. Optimal map-making is memory/IO-bound, pushing to massively-parallel HPC and GPU/exascale.
- **Systematics into C_N .** Gain, pointing, bandpass, half-wave plates, $T \rightarrow P$ leakage must be propagated into the noise model.
- **The covariance wall.** The dense pixel–pixel C_N is infeasible at native resolution — the same memory + noise-covariance limit as the exact likelihood.

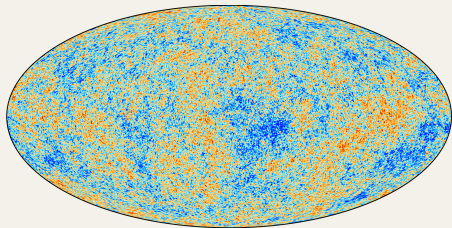
$D = \text{pixels}$ — the exact likelihood

PIXEL-BASED · THE GOLD STANDARD

$$-2\log\mathcal{L} = \mathbf{m}^\top \text{COV}^{-1} \mathbf{m} + \log|\text{COV}|$$

$$\mathbf{m} = \{T, Q, U\}, \quad \text{COV} = \mathbf{S}(C_\ell) + \mathbf{C}_N$$

- **Exact**: a multivariate Gaussian in *pixel* space, **even on the cut sky** — here $\mathbf{C}_N$ is the **dense pixel–pixel** map-making covariance.
- **Template systematics** marginalized analytically: add $\sigma^2 \mathbf{T} \mathbf{T}^\top$ to COV, $\sigma^2 \rightarrow \infty$ (foreground, monopole/dipole).
- Cost cubic, $\mathcal{O}(N_{\text{pix}}^3)$ — long held **intractable beyond very low resolution**.



Planck temperature map $T(\hat{n})$.

PICSLike — the exact likelihood, made easy

REMOVING THE HISTORICAL ENTRY BARRIER

- PICSLike is a **fast, user-friendly** engine for the exact pixel-based likelihood — a few lines of config, not a bespoke pipeline.
- It dissolves two historical barriers: the **entry cost** of writing one, and the **difficulty at large scales** — now **fully tractable** at realistic resolution.
- The bottleneck is no longer CPU but **memory** and the **noise-covariance model**.
- **Fully modular**: cleanly generalizes to **template fitting**, new noise models, extra components, and more.



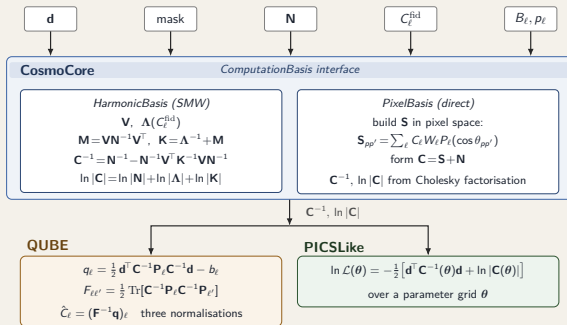
LiteBIRD-like regime

- $N_{\text{side}} = 16$, $f_{\text{sky}} \approx 0.5$: **interactive** ($\sim$ seconds).
- $N_{\text{side}} = 64$, $f_{\text{sky}} \approx 0.6$, joint Q , U : **~ 47 min** on one node.

Exact, end-to-end, at real survey sizes.

CosmoForge — one framework

ONE WORKSPACE: COSMOCORE · QUBE · PICSLIKE



`pip install cosmoforge · v1.0.1 · QML estimator + exact pixel likelihood, one config.`

$$D = C_\ell \text{ (full sky)}$$

SPECTRUM-BASED · WHY WE COMPRESS

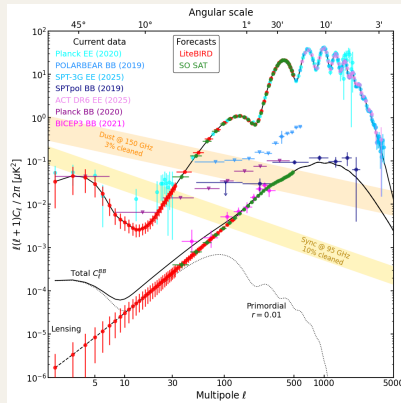
$$a_{\ell m} = \int Y_{\ell m}^*(\theta, \varphi) \frac{\Delta T}{T_0} d\cos\theta d\varphi$$

$$\langle a_{\ell m}^* a_{\ell' m'} \rangle = \delta_{\ell\ell'} \delta_{mm'} C_\ell$$

Wishart form (full sky):

$$-2 \log \mathcal{L} = \sum_{\ell} (2\ell + 1) \left[\frac{\hat{C}_\ell}{C_\ell} - \log \frac{\hat{C}_\ell}{C_\ell} - 1 \right]$$

- Compresses maps to angular spectra — cheap, and **exact on the full sky**; scales where pixel-based cannot.
- Some systematics are natural **here**: the **beam** is (nearly) diagonal in ℓ , and **$E \rightarrow B$ leakage** from the mask is handled by purification / pure- B estimators.



EE / BB spectra: current data and forecasts.

The catch — galactic foregrounds force a mask

COMPLICATION

Spherical harmonics are orthonormal and complete on the *full* sphere:

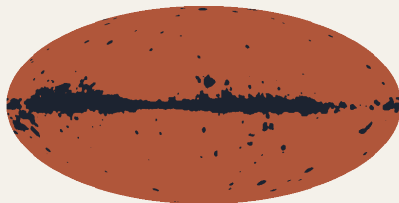
$$\int_{S^2} Y_{\ell m}^* Y_{\ell' m'} d\Omega = \delta_{\ell\ell'} \delta_{mm'}.$$

A mask restricts the integral to a subset of the sky:

$$\int_{S^2 \setminus \text{mask}} Y_{\ell m}^* Y_{\ell' m'} d\Omega = W_{mm'}^{\ell\ell'} \neq \delta_{\ell\ell'} \delta_{mm'}.$$

Different multipoles **couple** ($M_{\ell\ell'}$ broadens around the diagonal) and the Wishart form fails:

$$-2 \log \mathcal{L} \neq \sum_{\ell} (2\ell + 1) \left[\frac{\widehat{C}_{\ell}}{\bar{C}_{\ell}} - \log \frac{\widehat{C}_{\ell}}{\bar{C}_{\ell}} - 1 \right].$$



full sky: $\delta_{\ell\ell'}$



cut sky: banded

Two first attempts

APPROXIMATIONS 1 & 2 OF 4

Fiducial Gaussian

$$-2 \log \mathcal{L} = (\hat{\mathbf{C}}_{\text{obs}} - \mathbf{C}_{\text{th}}) \text{COV}_{\text{fid}}^{-1} (\hat{\mathbf{C}}_{\text{obs}} - \mathbf{C}_{\text{th}})^{\top}$$

~ full-sky exact

✓ $\ell\ell'$ couplings

✗ non-Gaussian C_{ℓ}

– negative- C_{ℓ} safe

- Couplings via a *fixed* covariance.
- Assumes Gaussian C_{ℓ} — fails at low ℓ .

Exact-with- f_{sky}

$$-2 \log \mathcal{L} = f_{\text{sky}} \sum_{\ell} (2\ell+1) \left[\frac{C_{\ell}^{\text{obs}}}{C_{\ell}^{\text{th}}} - \log \frac{C_{\ell}^{\text{obs}}}{C_{\ell}^{\text{th}}} - 1 \right]$$

✓ full-sky exact

✗ $\ell\ell'$ couplings

✓ non-Gaussian C_{ℓ}

– negative- C_{ℓ} safe

- Keeps the non-Gaussian shape; rescales the d.o.f. by the sky fraction f_{sky} .
- Disregards multipole couplings entirely.

Each fixes one failure of the other — neither captures couplings and non-Gaussianity at once.

Approximation 3 / 4 — Hamimeche–Lewis (HL)

APPROXIMATION 3 OF 4

✓ full-sky exact

✓ $\ell\ell'$ couplings

✓ non-Gaussian C_ℓ

✗ negative- C_ℓ safe

$$-2 \log \mathcal{L} = \left[\overrightarrow{g\left(\frac{c_{\text{obs}}}{c_{\text{th}}}\right) \mathbf{c}_{\text{fid}}} \right] \text{COV}_{\text{fid}}^{-1} \left[\overrightarrow{g\left(\frac{c_{\text{obs}}}{c_{\text{th}}}\right) \mathbf{c}_{\text{fid}}} \right]^\top$$
$$g(x) = \text{sign}(x - 1) \sqrt{2(x - \log x - 1)}$$

- $g(\cdot)$ symmetrises the Wishart shape around 1.
- Captures multipole couplings *and* non-Gaussianity.
- Negative C_ℓ values break the square root.

Approximation 4 / 4 — Offset HL

APPROXIMATION 4 OF 4

✓ full-sky exact

✓ $\ell\ell'$ couplings

✓ non-Gaussian C_ℓ

✓ negative- C_ℓ safe

$$-2 \log \mathcal{L} = \left[g\left(\frac{\mathbf{c}_{\text{obs}} + \mathbf{o}_\ell}{\mathbf{c}_{\text{th}} + \mathbf{o}_\ell}\right) (\mathbf{c}_{\text{fid}} + \mathbf{o}_\ell) \right] \text{COV}_{\text{fid}}^{-1} \left[g\left(\frac{\mathbf{c}_{\text{obs}} + \mathbf{o}_\ell}{\mathbf{c}_{\text{th}} + \mathbf{o}_\ell}\right) (\mathbf{c}_{\text{fid}} + \mathbf{o}_\ell) \right]^\top$$

$$g(x) = \text{sign}(x) \text{sign}(|x| - 1) \sqrt{2(|x| - \log|x| - 1)}$$

- Same machinery as HL with an additive offset $\mathbf{o}_\ell$.
- Offsets **shift away** from the negative- C_ℓ problem.
- Workhorse for low- ℓ B-mode polarisation.

From one field to many

THE QUESTION THAT GIVES BIRTH TO MHL

- The **spectrum** likelihoods so far — Fiducial Gaussian, exact- f_{sky} , HL, offset HL — were **single-field**: one map, one spectrum per ℓ .
- But real analyses — and the future — **combine many fields and experiments**: channels, splits, whole surveys of the same sky.
- The challenge: **consistency checks** *between* experiments and *within* one (as we'll do for BICEP). When noise biases can't be modelled perfectly, **cross-spectra may be the only robust way forward**.

Two questions

1. How do we combine **multiple fields** into *one* coherent likelihood?
2. And how do we use *only the cross-spectra* — the part we trust — **rigorously**?

Multi-field setup — and the noise-bias problem

THE STARTING POINT

We proceed with **HL** — the best-performing approximation at large scales. With n_{ch} fields (channels or splits) the spectra at each ℓ form a symmetric matrix (e.g. $n_{\text{ch}} = 3$):

$$\mathbf{C}_\ell = \begin{pmatrix} C_\ell^{A \times A} + N_\ell^A & C_\ell^{A \times B} & C_\ell^{A \times C} \\ C_\ell^{A \times B} & C_\ell^{B \times B} + N_\ell^B & C_\ell^{B \times C} \\ C_\ell^{A \times C} & C_\ell^{B \times C} & C_\ell^{C \times C} + N_\ell^C \end{pmatrix}$$

- **Auto**-spectra (diagonal) contain $\langle |\text{noise}|^2 \rangle$: we subtract an estimate $\hat{N}_\ell^X$. If $\hat{N}_\ell^X \neq N_\ell^{X, \text{true}}$, the data vector is **biased** $\Rightarrow$ parameters shift.
- **Cross**-spectra carry no noise bias in expectation — the **robust** part of the data vector $X_\ell = (X_\ell^{AA}, X_\ell^{AB}, \dots, X_\ell^{CC})$.

The fix — marginalized HL (mHL)

CROSS-ONLY BY MARGINALIZATION [NEW]

Keep **only the trusted crosses**. Empirically $X_\ell \sim \mathcal{N}(\mu, M_{\ell\ell'})$, so apply the standard Gaussian marginalization rule:

$$Y \sim \mathcal{N}(\mu, \Sigma) \implies Y_{\text{cross}} \sim \mathcal{N}(\mu_{\text{cross}}, \Sigma_{\text{cross, cross}})$$

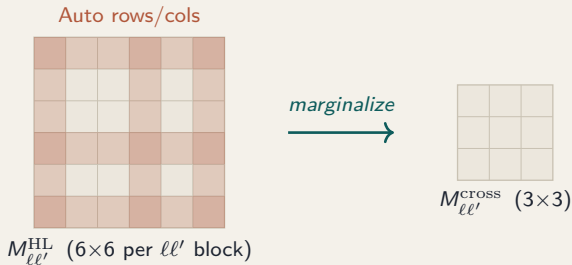
Build $X_\ell^{\text{cross}} = (X_\ell^{A \times B}, X_\ell^{A \times C}, X_\ell^{B \times C})$ and **drop the rows/cols of $M_{\ell\ell'}$ for the auto-elements**:

$$-2 \log \mathcal{L} = \sum_{\ell\ell'} X_\ell^{\text{cross} T} [M^{\text{cross}}]_{\ell\ell'}^{-1} X_{\ell'}^{\text{cross}}$$

mHL answers both questions — combine all fields, lean only on the trusted crosses.

The geometric picture — strip auto rows/cols

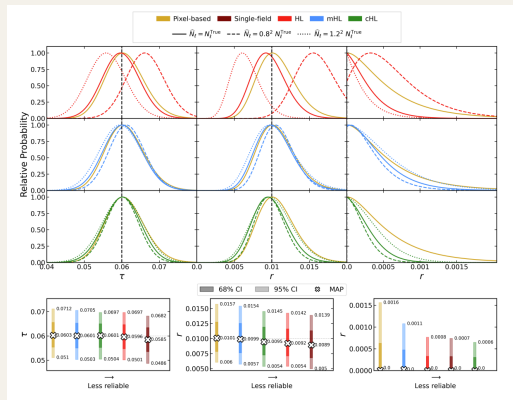
WHAT MHL DOES TO THE COVARIANCE



Drop the rows and columns associated with $X^{X \times X}$. No new offsets, no per-cross transformation — just a sub-block of the original covariance.

Validation — and the takeaway

CUT-SKY $f_{\text{sky}} = 40\%$ VS THE EXACT PIXEL-BASED TRUTH



Setup: 3 channels, $\ell_{\text{max}} = 32$, $\pm 20\%$ noise-bias error; ground truth = exact pixel-based.

- All HL approximations are **over-confident** vs PB — **mHL the least**.
- Exact within HL, no new offsets; **robust** to noise misestimation.
- Multi-field mHL **beats** single-field HL on the combined map.
- Public code: CHARM-Like.

Now: put it to work on real data.

MHL ON REAL DATA · IN BRIEF

- ELiCA combines Planck 100, 143 GHz and 70 GHz + WMAP as three fields in one mHL likelihood.
- It uses all the large-scale polarisation data available today, combined for the first time, to update the optical depth τ .

Genesini et al. 2026 — arXiv:2603.22454.

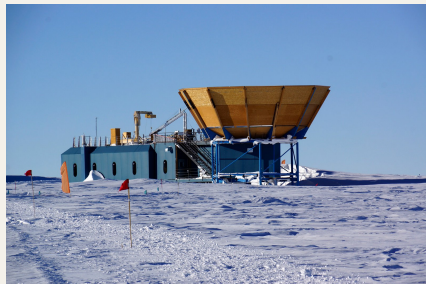
A whole talk of its own

Led by my student **V. Genesini** — the results are in her dedicated talk. Here, just proof that mHL works on real data.

Diagnosing BICEP/Keck 2018 with mHL

THE SETTING

- Constraints on the tensor-to-scalar ratio r from low- ℓ BB polarisation.
- 11 frequency maps combined into one likelihood:
 - BICEP3/Keck at 95, 150, 220 GHz + B95_e
 - Planck at 30, 44, 143, 217, 353 GHz
 - WMAP at 23 and 33 GHz
- **Eight sampled parameters:** $r + 7$ foreground (dust & synchrotron sector).
- Many auto-spectra with different noise properties
⇒ **natural setting for mHL.**



BICEP Array at the South Pole.

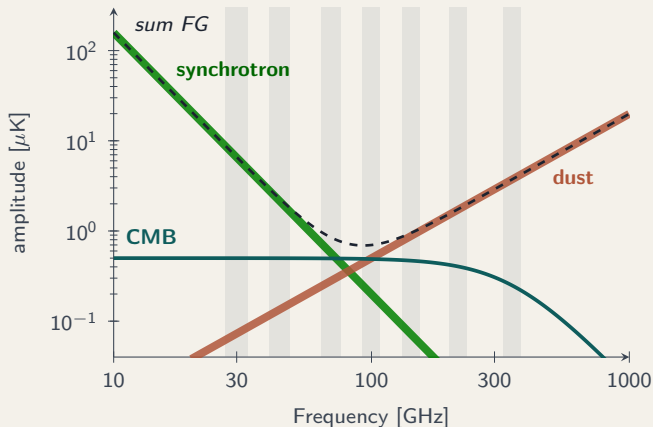
Galactic foregrounds — dust and synchrotron

WHY SO MANY CHANNELS AND PARAMETERS

The **same foregrounds** that forced the mask earlier now set the channel count. Two polarised galactic emissions dominate the large-scale B-mode sky:

- **Dust** — modified blackbody, rising at $\nu \gtrsim 100$ GHz ($\beta_d \sim 1.6$).
- **Synchrotron** — e^- in galactic **B**, falling at $\nu \lesssim 50$ GHz ($\beta_s \sim -3.1$).

Both smooth on the sky ($D_\ell^{\text{fg}} \propto \ell^\alpha$, $\alpha \sim -0.5$): only **frequency dependence** separates them from CMB — the job of **component separation** (see A. Carones, N. Krachmalnicoff).



The test — mHL as a consistency tool

MHL IS MORE THAN A ROBUST r

mHL rests on one fact: after the g -transform the data vector X is (empirically) **Gaussian**. So we apply the standard multivariate marginalization rule to keep **only the cross-spectra**.

- That already gives a **consistency check**: if the result *moves* when we marginalize a channel/auto, something may be **wrong with the autos**.
- Pushing the same Gaussianity further, the **conditional** likelihood is known too — so the full BICEP likelihood **factorizes exactly**, no approximation:

$$\mathcal{L}_{\text{full}}(\theta) = \underbrace{\mathcal{L}_{\text{marg}}(\theta)}_{\text{cross only (mHL)}} \cdot \underbrace{\mathcal{L}_{\text{cond}}(\theta)}_{\text{auto} \mid \text{cross}}$$

- This unlocks a **much deeper** test: **cross** (marginal) vs **auto | cross** (conditional) accordance, parameter by parameter.

Quantifying agreement — suspiciousness

A PRIOR-INDEPENDENT TENSION MEASURE

The Bayes ratio R measures concordance of two posteriors (here **cross** vs **auto**) but depends on the prior. The **information** I — how much the data shrink the prior — depends on it the *same way*, so the difference cancels it:

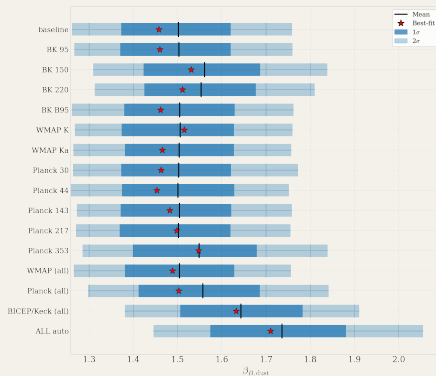
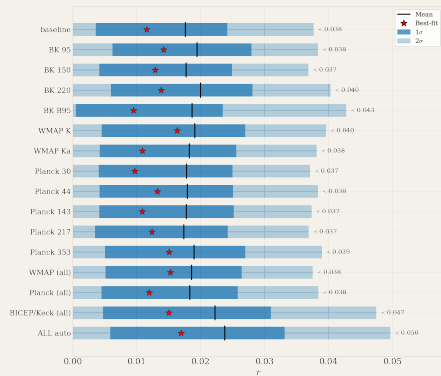
$$\ln R = \ln \frac{Z_{12}}{Z_1 Z_2}, \quad \ln I = D_{\text{KL},1} + D_{\text{KL},2} - D_{\text{KL},12}, \quad \boxed{\ln S = \ln R - \ln I}$$

- $\ln S$ (**suspiciousness**) is therefore **prior-independent**.
- $d - 2 \ln S \sim \chi_d^2$ (d = effective shared constrained parameters): it scores both a **discrepancy** ($\ln S > 0$) and an **over-concordance** ($\ln S < 0$); $\ln S = 0$ is perfect concordance.

Marshall et al. 2006 · Handley & Lemos 2019.

mHL on BK18 — marginalized constraints

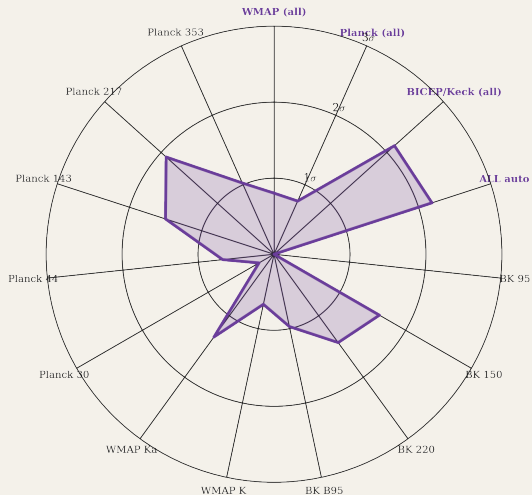
APPLY *JUST* MHL — MARGINALIZE EACH CHANNEL · PRELIMINARY



Baseline $r < 0.038 \rightarrow$ all-auto marginalized $r < 0.050$: the 95% bound *degrades by $\approx 32\%$* , while β_{dust} *stays put* — mHL alone, no conditional yet.

Suspiciousness across channels

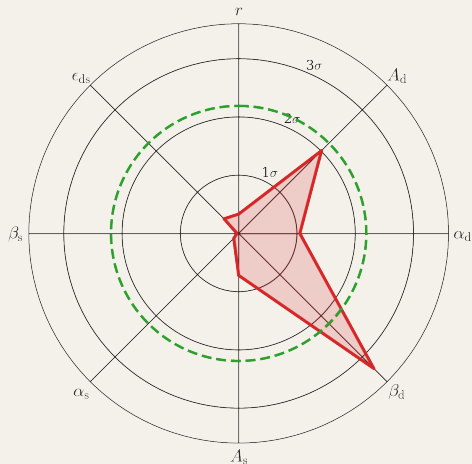
CROSS VS AUTO, PER CHANNEL AND COMBINATION · PRELIMINARY



- Suspiciousness between the **cross** (marginal) and **auto | cross** (conditional) posteriors, per channel.
- **All-auto** reaches 2.2σ ; all BICEP/Keck 2.1σ ; Planck / WMAP alone $\lesssim 0.8\sigma$.
- Single channels are weakly-constrained conditionals — not a shared-parameter signal.

Where the discrepancy lives — parameter shifts

PER PARAMETER, MARGINAL VS CONDITIONAL · PRELIMINARY

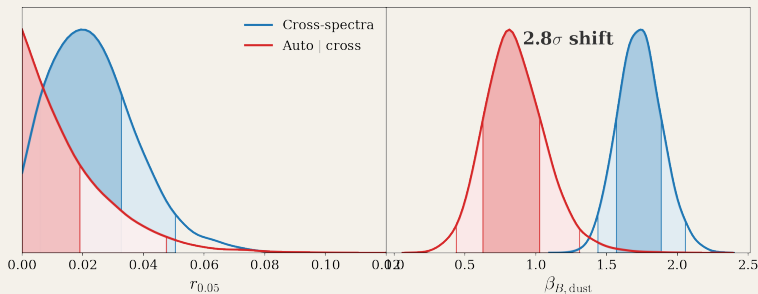


Shift between the **cross** (marginal) and **auto | cross** (conditional) means:

$$S_i = \frac{(\mu_i^{\text{cr}} - \mu_i^{\text{cd}})^2}{\sigma_i^{2,\text{cr}} + \sigma_i^{2,\text{cd}}}.$$

- “cd” is the **autos given the crosses** — *not* a standalone auto fit.
- β_{dust} and A_{dust} stand out — the **whole dust sector**; r and the rest stay small.

CROSS (MARGINAL) VS AUTO | CROSS (CONDITIONAL) · PRELIMINARY

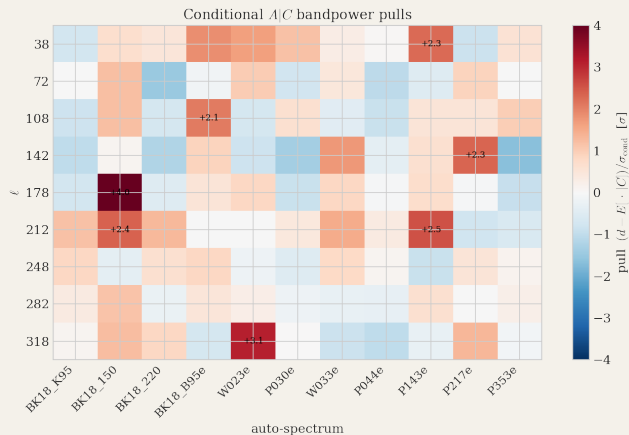


- $r_{0.05}$ **stable**: cross 0.024 vs auto 0.017 (0.3σ).
- β_{dust} **shifts**: cross 1.74 ± 0.16 vs auto $0.86 \pm 0.22 \Rightarrow 2.8\sigma$ (non-Gaussian, quoted — the naive Gaussian estimate gives 3.3σ).

The “auto” curve is the conditional (autos given the crosses), not a standalone auto fit.

Where in the data — bandpower pulls

AUTO | CROSS RESIDUALS, PER CHANNEL AND MULTIPOLE · PRELIMINARY



Standardized residual of each auto bandpower against what the **cross** predicts for it:

$$\text{pull}_\ell = \frac{\hat{X}_\ell^a - \bar{X}_\ell^a}{\sigma_{a|c}}.$$

- Dominated by **one cell**: BK 150 GHz at $\ell \approx 178$, $+4.04\sigma$.
- No other cell exceeds $\sim 3.1\sigma$ (a lone noisy WMAP 23 GHz bin).
- Not broadband — **one channel**, **one multipole bin**.

All BICEP results preliminary — Galloni & Ballardini, in prep.

A diagnostic tool

FROM A LIKELIHOOD FIX TO A DIAGNOSTIC TOOL

- mHL's exact factorization is a **powerful, model-faithful diagnostic** of an experiment's **self-consistency**: here, r robust, a dust-sector cross/auto discrepancy.
- Naturally **promising for multi-experiment** setups — the same construction cross-checks *different experiments*.
- All data points are *first-class citizens* of one data vector (like different fields), so marginalizing the **high- ℓ** part **pinpoints the low- ℓ** origin of the discrepancy.
- The **origin is an open question** — a residual effect in the deep BK auto-spectra; we do *not* present it as a detection.

All BICEP results preliminary — Galloni & Ballardini, in prep.

Open challenges — and two we softened

IN ONE SLIDE

- Extracting parameters at **large scales** is hard — the likelihood is non-trivial at every step (map-making, the exact pixel likelihood, the C_ℓ approximations), and the open challenges are **many**.
- Today we softened **two**:
 - PICSLike — the exact pixel likelihood, now fast and tractable at large scales;
 - mHL — robust multi-field C_ℓ likelihoods, *and* an exact-factorization self-consistency test.
- On real data: r robust from BICEP/Keck (with a preliminary dust-sector cross/auto discrepancy), and τ from all large-scale polarisation (V. Genesini).
- The same direction: making **future, harder** CMB measurements — down to B-modes at $r \sim 10^{-3}$ — trustworthy.

Thank you

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APPENDIX

Backup slides

Material for questions.

HL likelihood — the eigenvalue trick

BACKUP · HAMIMECHE & LEWIS 2008

Start from the exact full-sky Wishart likelihood at fixed ℓ , with $\hat{\mathbf{C}}_\ell \sim \mathcal{W}_n(\mathbf{C}_\ell, 2\ell + 1)$.

Diagonalize:

$$\mathbf{D}_\ell = \mathbf{C}_\ell^{-1/2} \hat{\mathbf{C}}_\ell \mathbf{C}_\ell^{-1/2} = \mathbf{U}_\ell \mathbf{\Lambda}_\ell \mathbf{U}_\ell^T,$$

and define the change of variable

$$\mathbf{g}(\mathbf{D}_\ell) = \mathbf{U}_\ell \text{diag}[g(\lambda_i)] \mathbf{U}_\ell^T, \quad g(x) = \text{sign}(x - 1) \sqrt{2(x - \log x - 1)}.$$

In this basis, the data vector is approximately Gaussian:

$$-2 \log \mathcal{L}_{\text{HL}} = \sum_{\ell\ell'} [\mathbf{g} \mathbf{C}_\ell^{\text{fid}}]_\ell^T M_{\ell\ell'}^{-1} [\mathbf{g} \mathbf{C}_{\ell'}^{\text{fid}}]_{\ell'},$$

with $M_{\ell\ell'}$ estimated from simulations. Cut sky / coupling: M also picks up the off-diagonal $\ell\ell'$ correlations.

mHL — Schur complement of the partitioned Gaussian

BACKUP · WHY MHL IS EXACT (WITHIN HL)

Partition the data vector into auto and cross blocks, $X_\ell = (X_\ell^a, X_\ell^c)$, with $M = \begin{pmatrix} M_{aa} & M_{ac} \\ M_{ca} & M_{cc} \end{pmatrix}$.

Marginal (drop autos):

$$p(X^c) \propto \exp\left[-\frac{1}{2}(X^c)^T M_{cc}^{-1} X^c\right] \implies \mathcal{L}_{\text{marg}} \text{ uses } M_{cc}.$$

Conditional (autos given crosses):

$$p(X^a | X^c) \propto \exp\left[-\frac{1}{2}(X^a - \bar{X}^a)^T \Sigma_{a|c}^{-1} (X^a - \bar{X}^a)\right],$$

with $\bar{X}^a = M_{ac} M_{cc}^{-1} X^c$ and $\Sigma_{a|c} = M_{aa} - M_{ac} M_{cc}^{-1} M_{ca}$ (Schur complement).

Factorization: $\mathcal{L}_{\text{full}}(\theta) = \mathcal{L}_{\text{marg}}(\theta) \cdot \mathcal{L}_{\text{cond}}(\theta)$ — exact within HL, enables the auto-vs-cross diagnostic of Part V.

Sellentin–Heavens correction

BACKUP · FINITE- N_{sim} COVARIANCE

$M_{\ell\ell'}$ is estimated from N_{sim} realisations $\Rightarrow$ it is a noisy Wishart-distributed matrix.

Marginalising over the true covariance gives a t -distribution-like form:

$$-2 \log \mathcal{L} = N_{\text{sim}} \log \left[1 + \frac{1}{N_{\text{sim}} - 1} X^T \hat{M}^{-1} X \right].$$

Reduces to the Gaussian when $N_{\text{sim}} \rightarrow \infty$; prevents posterior over-confidence for small N_{sim} .

In ELiCA: $N_{\text{sim}} = 500$, $n_{\text{data}} \approx 28 \times 6 = 168$ — correction is non-negligible.

Sellentin & Heavens 2015 — arXiv:1511.05969

In R , In I , In S — formulas

BACKUP · SUSPICIOUSNESS

Bayes ratio for two datasets D_1, D_2 on a shared parameter θ :

$$R = \frac{Z_{12}}{Z_1 Z_2} = \frac{P(D_1, D_2)}{P(D_1) P(D_2)},$$

with evidences $Z_i = \int \mathcal{L}_i(\theta) \pi(\theta) d\theta$. Information from likelihoods:

$$D_{\text{KL},i} = \int P_i(\theta) \log \frac{P_i(\theta)}{\pi(\theta)} d\theta, \quad \ln I = D_{\text{KL},1} + D_{\text{KL},2} - D_{\text{KL},12}.$$

Suspiciousness (prior-independent):

$\ln S = \ln R - \ln I$

Both $\ln R$ and $\ln I$ are estimated from nested sampling along with the evidences.

Marshall et al. 2006 · Handley & Lemos 2019

Calibration — $d - 2 \ln S \sim \chi_d^2$

BACKUP · TURNING $\ln S$ INTO A TENSION SCALE

For Gaussian posteriors of dimension d , the suspiciousness is calibrated by

$$d - 2 \ln S \sim \chi_d^2.$$

d = effective number of constrained shared parameters (Bayesian model dim.):

$$d = 2 \left(\langle (\ln \mathcal{L})^2 \rangle_P - \langle \ln \mathcal{L} \rangle_P^2 \right).$$

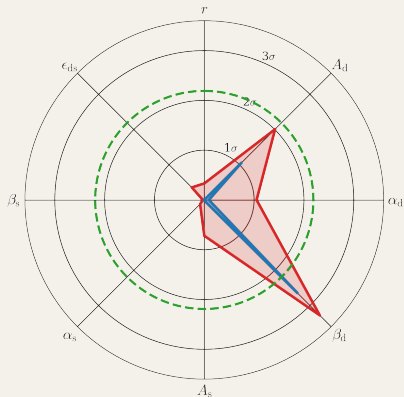
Tension probability:

$$p = \int_{d-2 \ln S}^{\infty} \chi_d^2(x) dx, \quad n_\sigma = \sqrt{2} \operatorname{erfc}^{-1}(p).$$

Why the dimensionality: a high-dim. shift can have small T per axis but a large total χ_d^2 — averaging across d dilutes signals (slide 46).

The look-elsewhere effect (LEE)

BACKUP · CORRECTING THE PER-PARAMETER SHIFTS



Scanning **8 parameters** inflates the largest shift — the **look-elsewhere effect**; the trials factor folds it in.

- **LEE-corrected** here (vs uncorrected in the main flow): the ring moves, β_{dust} still stands out.
- Localizes to a single deep BK **150 GHz** bandpower ($\ell \approx 178$); overall significance spans $\approx 2.2\text{--}2.8\sigma$ — reduced, not erased.

Preliminary.

Is it just that one bandpower?

BACKUP · ROBUSTNESS CHECKS · PRELIMINARY

- Drop the peak bandpower (BK 150 GHz, $\ell \approx 178$) and repeat the full fit:
 - Conditional β_{dust} moves back toward the cross, $0.86 \pm 0.22 \rightarrow 1.01 \pm 0.28$ (cross unchanged, 1.74 ± 0.16).
 - Shift drops $2.8\sigma \rightarrow 2.1\sigma$; omnibus suspiciousness $2.2\sigma \rightarrow 1.7\sigma$.
 - r unaffected — the feature lives in the foreground sector, not the cosmology.
- So the single cell carries about half the effect; a coherent residual across the remaining BK auto-bandpowers carries the other half — localized, not pointlike.
- An untargeted omnibus χ^2 over all 99 conditional residuals is unremarkable (0.9σ): the discrepancy is coherent and localized, not a broadband excess.

All BICEP results preliminary — Galli & Ballardini, in prep.

The non-Gaussian parameter-shift test

BACKUP · THE ACCURATE β_{dust} SHIFT

The Gaussian shift assumes both posteriors are Gaussian. The **conditional** β_{dust} is skewed and pressed against its lower bound (0.86 with a tail toward 0), so the Gaussian σ **overstates** the separation. Form $\Delta\beta = \beta_{\text{cross}} - \beta_{\text{auto}}$ from the two independent posteriors; the tension is the chance there *is* a shift:

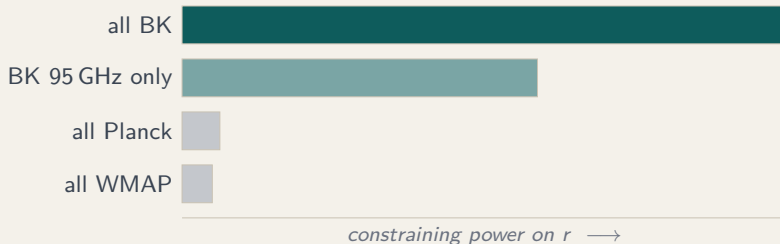
$$P_{\text{shift}} = \int_{p(\Delta\beta) > p(0)} p(\Delta\beta) d\Delta\beta, \quad n_\sigma = \sqrt{2} \operatorname{erf}^{-1} P_{\text{shift}}.$$

- *Gaussian* limit: $n_\sigma = |\Delta\mu| / \sqrt{\sigma_c^2 + \sigma_a^2} = 3.3\sigma$.
- *Non-Gaussian* (KDE from the samples, no Gaussian assumption): $P_{\text{shift}} = 0.9952 \Rightarrow 2.82\sigma$ (stable 2.82/2.89 across resamplings) $\Rightarrow \approx 2.8\sigma$.
- All other 7 parameters $< 2\sigma$ by the same test — β_{dust} -localized.

Where does the r information actually come from?

BACKUP · CHANNEL-BY-CHANNEL

Restrict the analysis to one channel block at a time:



BK channels carry r ; Planck/WMAP enter as foreground templates. Removing all BK gives $r = 0.68 \pm 0.34$.

Map-making — GLS and destriping

BACKUP · TOD → MAP, IN DETAIL

The Gaussian TOD likelihood for $\mathbf{d} = \mathbf{P}\mathbf{m} + \mathbf{n}$ is $-2 \log \mathcal{L} = (\mathbf{d} - \mathbf{P}\mathbf{m})^\top \mathbf{N}^{-1}(\mathbf{d} - \mathbf{P}\mathbf{m})$. Setting $\partial/\partial\mathbf{m} = 0$ gives the normal equations

$$(\mathbf{P}^\top \mathbf{N}^{-1} \mathbf{P}) \hat{\mathbf{m}} = \mathbf{P}^\top \mathbf{N}^{-1} \mathbf{d}.$$

- $\mathbf{N}^{-1}$ acts as a **Fourier-domain noise filter** (stationary $1/f$); the system is solved iteratively by **preconditioned conjugate gradient**.
- **Destriping** splits the noise $\mathbf{n} = \mathbf{F}\mathbf{a} + \mathbf{w}$ into baseline offsets $\mathbf{F}\mathbf{a}$ (the $1/f$ part) and white $\mathbf{w}$. ML for the amplitudes,

$$\hat{\mathbf{a}} = (\mathbf{F}^\top \mathbf{Z} \mathbf{F} + \mathbf{C}_a^{-1})^{-1} \mathbf{F}^\top \mathbf{Z} \mathbf{d}, \quad \mathbf{Z} = \mathbf{C}_w^{-1} - \mathbf{C}_w^{-1} \mathbf{P} (\mathbf{P}^\top \mathbf{C}_w^{-1} \mathbf{P})^{-1} \mathbf{P}^\top \mathbf{C}_w^{-1},$$

then bin the cleaned stream. A baseline prior $\mathbf{C}_a$ regularizes the fit; the solution $\rightarrow$ full ML as the baseline length $\rightarrow$ one sample.

Keskitalo et al. 2005 (arXiv:astro-ph/0412517) · Biquard 2024.

Map-making — transfer functions and $\mathbf{C}_N$

BACKUP · WHY LARGE SCALES AND MEMORY ARE HARD

- **Filter-and-bin.** A time-domain filter applied before binning **biases the map**; the bias is calibrated as an empirical *transfer function* from input-vs-filtered simulations (CLASS: a pixel-space transfer matrix). Cheap, but the large-scale correction is the delicate part.
- **The covariance wall.** The dense **pixel-pixel** $\mathbf{C}_N = (\mathbf{P}^\top \mathbf{N}^{-1} \mathbf{P})^{-1}$ is $\mathcal{O}(N_{\text{pix}}^2)$ to store — infeasible at native resolution, and reconstructed only on **low-resolution** maps via Monte-Carlo / destripping-baseline estimation.
- **Scale.** Optimal map-making is memory/IO-bound, pushing to massively-parallel HPC and GPU/exascale.