

# Basic Concepts of Statistics

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**School Of Statistics**

Carry-le-Rouet – 16/05/2022

# General introduction

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## Goals of the lecture

- recap the **basics** needed for the SOS
- learn **how** to be **critical** with statistics (in science, but not only)
- *focus* on **meaning** and **(mis)intuition** rather than mathematical rigour

## **Statistics versus probability** (according to Persi Diaconis)

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## **Few personal tips for this lecture**

- keywords/concepts will be listed at the end of each section  
→ make sure you know the ideas behind them!

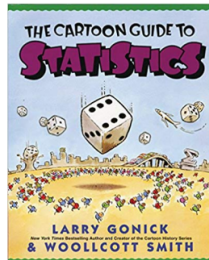
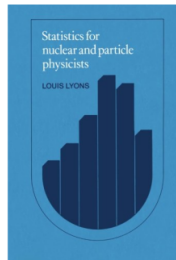
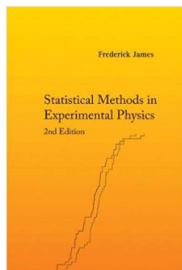
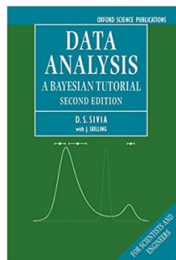
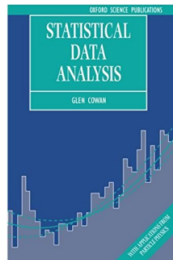
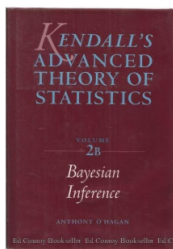
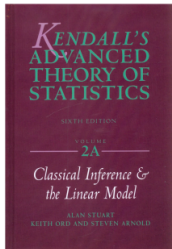
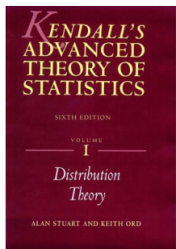
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- statistics is almost like a language: you need practice to learn it!  
→ compute/code as much as simple examples as you can **by yourself!**

# Some references



1. **Statistics**
2. **Probability**
3. **Statistical model**
4. **The two big schools**
5. **Parameter estimation and hypothesis testing**

# Statistics

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## Definitions:

- Descriptive statistics  $\sim$  “summarize” a sample
- sample = set of observations  $\mathcal{S} \equiv \{x_1, x_2, \dots, x_n\}$

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  - Skewness:  $\gamma_x = \overline{\left(\frac{x - \bar{x}}{\sigma_x}\right)^3}$  - asymmetry

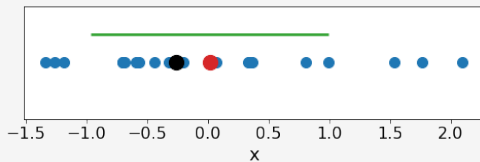
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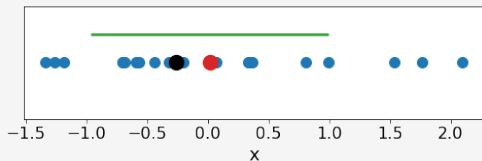
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  - Kurtosis:  $\beta_x = \left(\frac{x - \bar{x}}{\sigma_x}\right)^4$  - importance of tails

# Sample characterisation - illustrations



blue:  $x_j$ , red: mean. black: median, green:  $\sigma_x$

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## Skewness and Kurtosis (using probability functions)



Right plot: Kurtosis  $\gamma = \infty$  (red), 2 (blue), 1, 1/2, 1/4, 1/8, and 1/16 (gray), 0 (black)

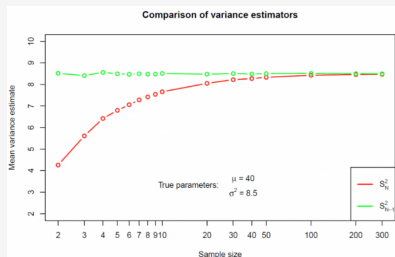
## **Notion of estimator** (more on this later)

- e.g.: sample mean  $\neq$  “true mean”
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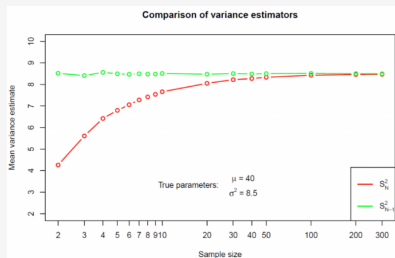
→ sample variance  $v_x$  is a **biased estimator** of the true variance.

But  $\frac{1}{n-1} \sum (x_i - \bar{x})^2$  is unbiased.

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## Statistical moments (more on this later)

- Order- $r$  moment:  $m_r = \overline{\left(\frac{x - \bar{x}}{\sigma_x}\right)^r}$  (relates directly to the mean of  $x^r$ )
- probability theory: all truth moments  $\equiv$  exact underlying probability
- **first** moments  $\equiv$  “**main**” features of the sample

## Multidimensional sample

- single observation  $i$  = several numbers:  $x_i \rightarrow (x_i^{(1)}, x_i^{(2)}, \dots, x_i^{(p)})$
- e.g. biological dataset: person size, weight, age and genre

Previous description applies to each variable  $x_i^{(j)}$  but one can **now** explore **how variables behave wrt each other**.



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**Covariance and correlations** between two variables  $a$  and  $b$ :

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- probes if fluctuations around the mean are coherent for  $a$  and  $b$
- covariance (and correlation) are symmetric - fortunate
- covariance of  $x$  with itself is the variance
- $\rho_{a,b} \in [-1, 1]$ ; 0 = uncorrelated ( $\neq$  indep!), (-)1 = (anti-)correlated

### **Covariance matrix** or error matrix

- $C_{ij} = \rho_{ij} \times \sigma_i \sigma_j$  - real and symmetric.
- $\rho_{ij}$  is the correlation matrix - symmetric with 1's on diagonal.

### **Why is this object so important?**

- find pattern in a dataset (e.g. is age correlated to weight?)

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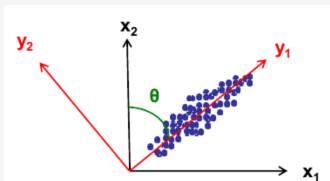
# More on correlations

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  - find directions which are uncorrelated (**Principal Component Analysis**)



- $x_1$  and  $x_2$  both have a large  $\sigma$
- but, they are highly correlated
- most of the information is in  $y_1$  (largest  $\sigma$ )
  - idea of dimension reduction
  - idea of pre-processing in ML

# Correlation and dependence

Correlation  $\equiv$  *linear* dependence  $\Rightarrow$  dependence

**BUT**

Non-correlation *doesn't* imply independence (matter of vocabulary)

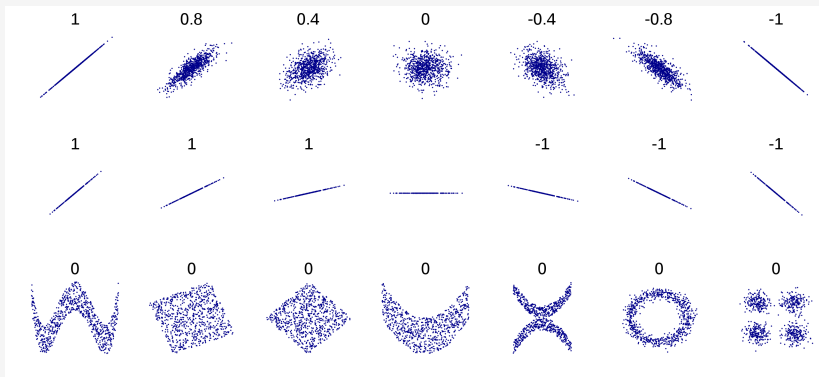


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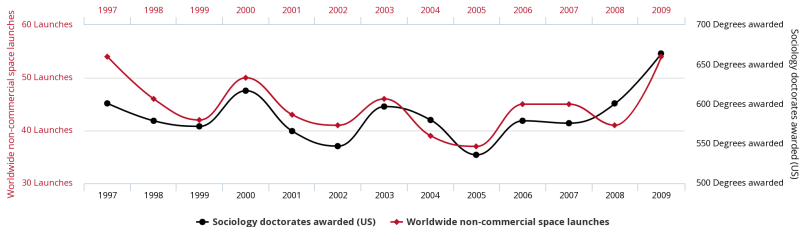
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**Worldwide non-commercial space launches**

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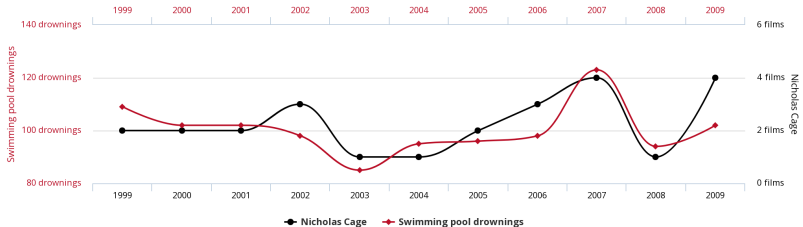


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**Number of people who drowned by falling into a pool**  
correlates with  
**Films Nicolas Cage appeared in**



tylervigen.com

## Part I

descriptive statistics – sample – mean – (co)variance – (de)correlation

# Probability

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Caution: what follows is *not* mathematically rigorous

## Random variable and associated probability

- a random variable  $X$  describes an **observable which is not certain**
- all possible outcomes - **realisations** - of  $X$  form a set  $\Omega$
- a probability  $P_i$  is associated to each realisation  $i$  of  $\Omega$
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## E.g. of the flipping coin

- $\mu = 1/2$ ,  $\sigma = 1/2$ ,  $m_r = 1$  if  $r$  is even and 0 if  $r$  is odd



### Bias theorem - math version

$$P(A|B) = P(A) \times \frac{P(B|A)}{P(B)}$$

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## Comments

- many ways to *understand* this fundamental equation
- in some case, each of these term has a [clear meaning](#)
- these two posts are quit interesting [post 1](#) and [post 2](#)

# Understanding Bias theorem

**Example:** *hypothesis = fire and evidence = smoke*

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→ this our **prior** knowledge about the hypothesis (often **arbitrary**)
- $P(\text{evidence}|\text{hypothesis})$ : proba that there is smoke if there is fire  
→ **easy to assess** (fire produces smoke)  
→ That is **the** interest of bias theorem



# Understanding Bias theorem

**Example:** *hypothesis = fire* and *evidence = smoke*

$$P(\text{fire}|\text{smoke}) = P(\text{fire}) \times \frac{P(\text{smoke}|\text{fire})}{P(\text{smoke})}$$

- $P(\text{hypothesis}|\text{evidence})$ : proba that there is a fire if there is smoke  
→ difficult to assess (many sources of smoke), that's **the posterior**
- $P(\text{hypothesis})$ : proba that there is a fire  
→ this our **prior** knowledge about the hypothesis (often **arbitrary**)
- $P(\text{evidence}|\text{hypothesis})$ : proba that there is smoke if there is fire  
→ **easy to assess** (fire produces smoke)  
→ That is **the** interest of bias theorem
- $P(\text{evidence})$ : proba that there is smoke somewhere  
→ the evidence is rare (valuable) to observe or not (indifferent)

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*N.B.:*  $P(\text{evidence})$  is independent from the hypothesis, and is sometime impossible to compute. It is often seen as a “normalization factor” and dropped while **comparing** different hypothesis.

## Few examples:

- I'm not feeling so well → Am I sick ?
- There are clouds → will it rain?
- I go out in a bar → will I end up drunk?
- I attend to a school statistics → will I learn something?

# Everyday life questions are often bayesian

## Few examples:

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## Always the same thinking:

1. you observe a fact
2. you wonder the probability of something, given you this fact happened
3. you have (sometimes rough/wrong) prior, based on past knowledge
4. your brain applies Bayes theorem, even you don't know it!

## Generalization to the continuous case

- There is a **whole continuum** of outcome (realization) for  $X$
- Probability described by a **density probability function (PDF)**,  $f(x)$ :

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- moments **are** the Taylor expansion coefficients:  $m_r = (-i)^r \left. \frac{d^r \varphi_x}{dt^r} \right|_{t=0}$

**Binomial law:** efficiency, trigger rates, ...

$$B(k; n, p) = C_k^n p^k (1-p)^{n-k}, \mu = np, \sigma = \sqrt{np(1-p)}$$

**Poisson distribution:** counting experiments, hypothesis testing

$$P(n; \lambda) = \frac{\lambda^n e^{-\lambda}}{n!}, \mu = \lambda, \sigma = \sqrt{\lambda}$$

**Gauss distribution (aka Normal):** many use-case (asymptotic convergence)

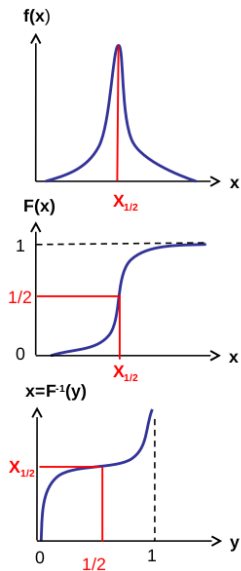
$$f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

**Cauchy distribution (aka Breit-Wigner):** particle decay width, ....

$$f(x; x_0, \gamma) = \frac{1}{\pi\gamma \left[ 1 + \left( \frac{x-x_0}{\gamma} \right)^2 \right]}$$

$\mu$  and  $\sigma$  not defined (divergent integral)

# Cumulative distribution and quantiles



Probability density function:  $f(x)$

Cumulative distribution:  $F(x)=y$

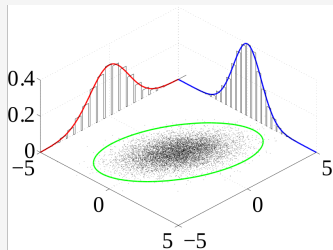
Inverse cumulative distribution:  $x=F^{-1}(y)$

**Median:**  $x$  such that  $F(x)=1/2 \rightarrow x_{1/2} = F^{-1}(1/2)$

**Quantile** of order  $\alpha$ :  $x_{\alpha} = F^{-1}(\alpha)$

## How to describe several random variables simultaneously?

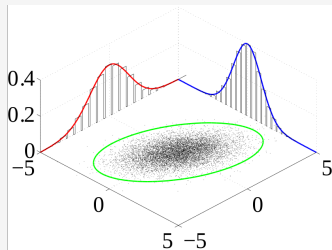
- $X$  and  $Y$  are two random variables  $\rightarrow$  PDF is  $f_{XY}$ ,
- **several questions** can be asked about  $X$ ,  $Y$  or both.



- Probability that  $X \in [x, x + dx]$  and  $Y \in [y, y + dy]$ :  
$$d^2P(x, y) = f_{XY}(x, y) dx dy$$
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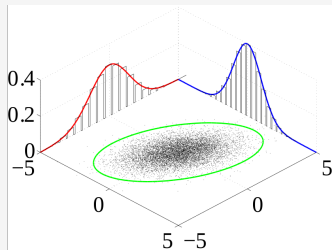
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## Multidimensional normal distribution

$$f(\vec{x}; \vec{\mu}, \Sigma) = \frac{1}{\sqrt{(2\pi)^n \det \Sigma}} \exp\left(-\frac{1}{2} (\vec{x} - \vec{\mu})^T \Sigma^{-1} (\vec{x} - \vec{\mu})\right)$$

- $\vec{\mu}$  mean position of  $\vec{x}$ ,  $\Sigma$  covariance matrix



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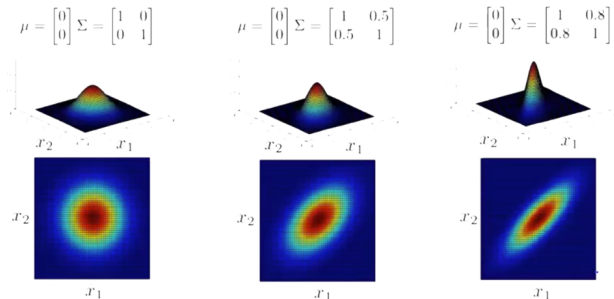
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# Central limit theorem

Caution: what follows is *not* mathematically rigorous

If  $n$  random variables  $\{X_i\}$  are distributed according to the same PDF  $f_X$  with a defined mean  $\mu_x$  and a std  $\sigma_x$ , then the random variable  $Y = \frac{1}{n}(X_1 + \dots + X_n)$  is following a normal distribution of mean  $\mu_x$  and std  $\sigma_x/\sqrt{n}$ .

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**For 2 variables**  $Y = X_1 + X_2$

- The PDF of  $Y$  is  $f_Y(y) = \int f_{X_1}(x_1) \times f_{X_2}(y - x_1) dx_1 \rightarrow$  convolution!
- Characteristic function:  $\varphi_Y(t) = \varphi_{X_1}(t) \times \varphi_{X_2}(t) = \varphi_X(t)^2$  - same PDF!
- 1st and 2nd moments known :  $\varphi_X(t) \sim$  2nd order Taylor expansion

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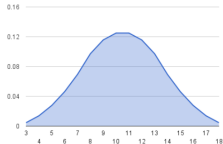
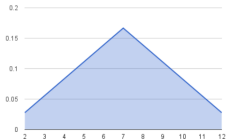
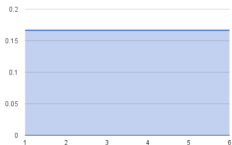
**Generalizing for sum of  $n$  variables:**

- $\varphi_Y(t) = \varphi_X(t)^n \sim \left(1 - \frac{t^2}{n}\right)^n \rightarrow e^{-t^2/2}$  for  $n \rightarrow \infty$
- going back to real space, a normal distribution is obtained

*N.B.* this reasoning doesn't explain why  $\sigma_Y = \sigma_X/\sqrt{n}$ , this needs to properly re-scale  $Y$ .

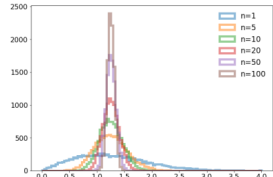
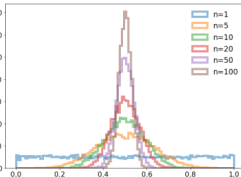
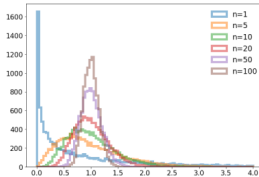
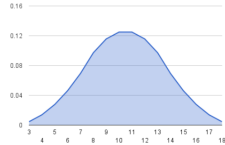
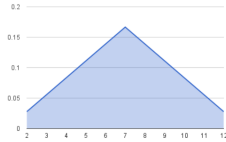
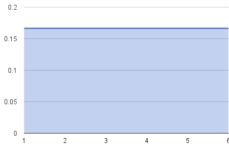
# Central limit theorem – continued

## One way to understand why it works



# Central limit theorem – continued

## One way to understand why it works



## Proof

Prove that  $\sigma_Y = \sigma_X/\sqrt{n}$  with the proper scalings to define  $Y$ .

## Application

Prove, using the CLT, that a Poisson distribution  $P(n; \lambda)$  tends to a normal distribution for large numbers.

*Hint:*  $N = 1 + 1 + 1 \dots + 1$  N-times



**Final observable** is very often a combination of (random) variable.

- $\mathcal{O} = g(X_1, X_2, \dots, X_n) \equiv g(\vec{X})$ .  $\mathcal{O}$  is also a **random variable**
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**Comments:**

- these equations are known as **error propagation**
- **this procedure is not exact** and relies on **Taylor expansion**
- only 1st and 2nd moments of  $\vec{X}$  are needed (or their **estimators**)

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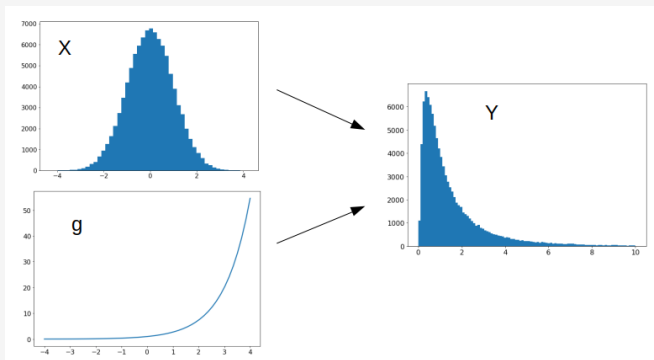
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## **Part I: statistics**

descriptive statistics – sample – mean – (co)variance – (de)correlation

## **Part II: probability**

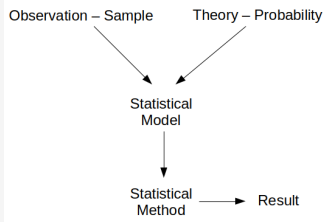
Bias theorem – prior – posterior – random variable – (marginal) PDF –  
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1. **Statistics**
2. **Probability**
3. **Statistical model**
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5. **Parameter estimation and hypothesis testing**

# Statistical model

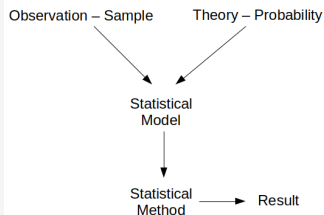
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# Statistical model: what, why, how



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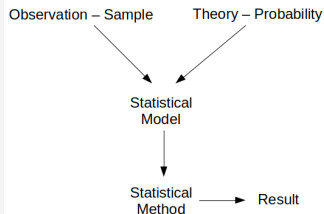
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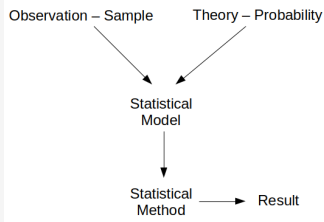
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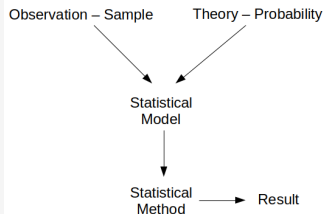
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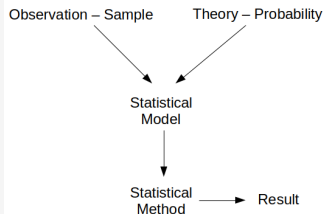
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A **statistical model** is also called **likelihood function**  $\mathcal{L}(\vec{\mu}, \vec{\theta}; \vec{x})$ . It can be seen as the **probability** that the physical model predicts the observable  $\vec{x}$ , **given the parameters**  $(\vec{\mu}, \vec{\theta})$ .

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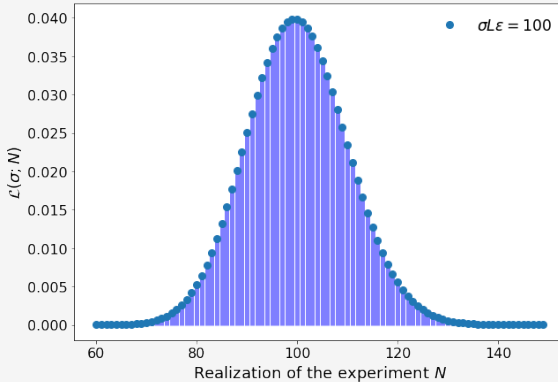
## Statistical model

$$\mathcal{L}(\sigma; N) = e^{-\sigma L \epsilon} \frac{(\sigma L \epsilon)^N}{N!}$$



# Illustration of the Likelihood

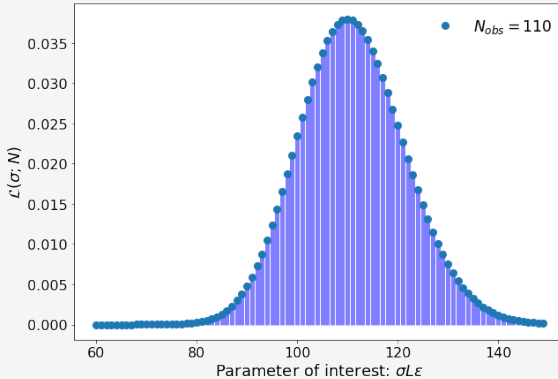
Given a value of  $\sigma$ , what's the “probability” to observe  $N$  ?



Anticipation: frequentist “usage” of the likelihood

# Illustration of the Likelihood

If we observed a value for  $N$ , what's the “probability” that  $\sigma = X$ ?



Anticipation: bayesian “usage” of the likelihood

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Systematic uncertainty estimation *and* treatment is **not** an exact science.

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# A first discussion on uncertainties

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## Implications:

- **arbitrariness** (and a looooot of discussion that go with it)
- always check the robustness of the conclusion wrt to those
- that's the way it is, no choice! → be *smartly* practical!



## **Part I: statistics**

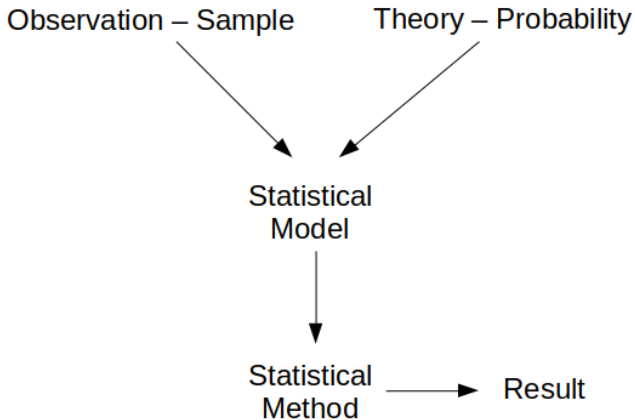
descriptive statistics – sample – mean – (co)variance – (de)correlation

## **Part II: probability**

Bias theorem – prior – posterior – random variable – (marginal) PDF –  
moments – characteristic function – (in)dependent variables –  
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## **Part III: statistical model**

Likelihood – nuisance parameter – parameter of interest –  
systematic uncertainties



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## The two big schools

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# Frequentist versus bayesian

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- *bayesian*: exploits the **Bayes theorem** to compute the posterior  $P(\text{para}|\text{obs})$ , using the **prior**  $P(\text{para})$  and  $P(\text{obs}|\text{para})$  - the **likelihood**

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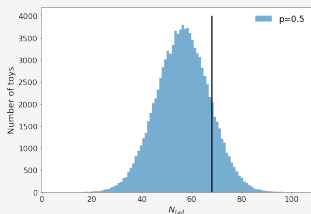
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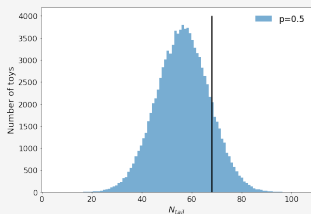
Let's try to analyze the **same experiment** with both **frequentist** and **bayesian** approaches

# Is a flipping coin tricked? Frequentist approach

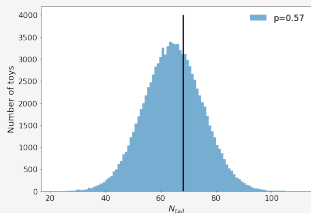


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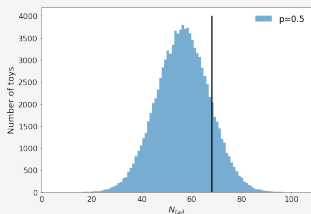


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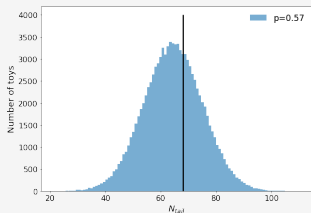


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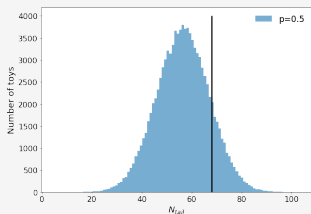


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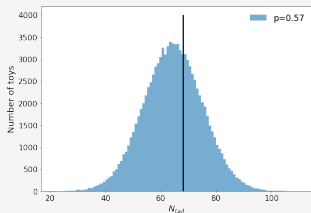
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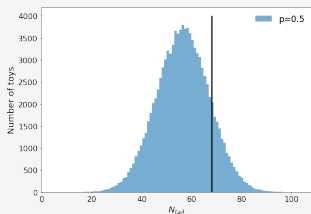


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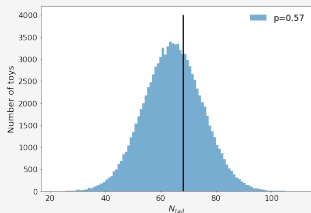
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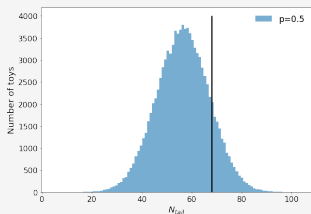


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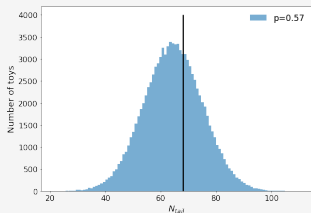
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- according to you, is  $p = 0.57$  more probable than  $p = 0.50$ ?  
→ this question has no sense in frequentist

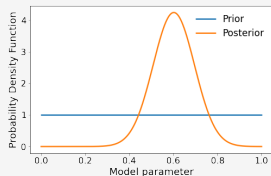
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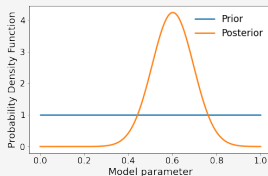
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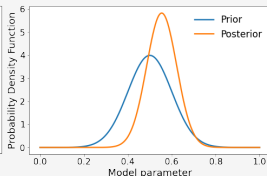
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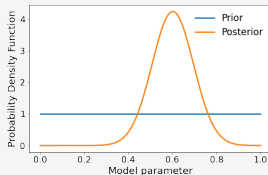


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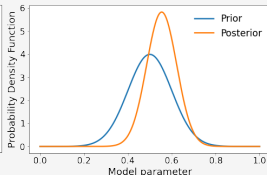
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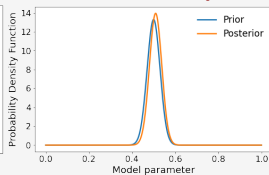
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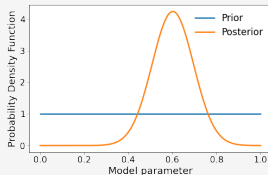


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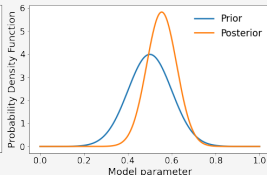
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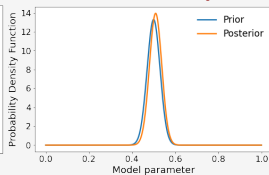
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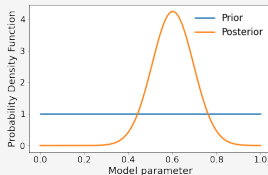
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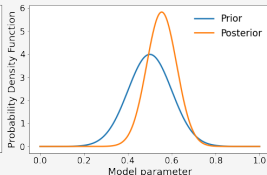
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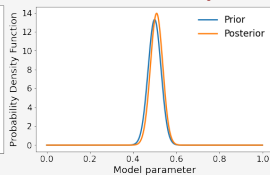
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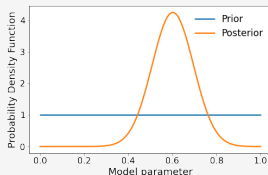
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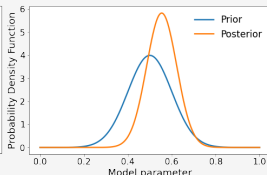
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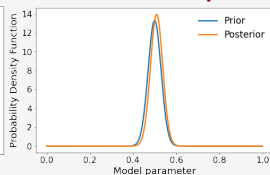
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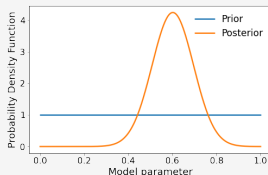
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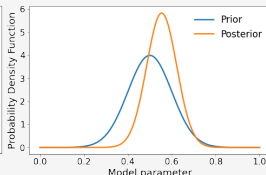
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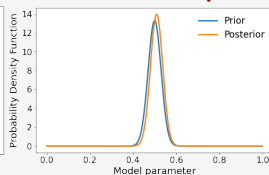
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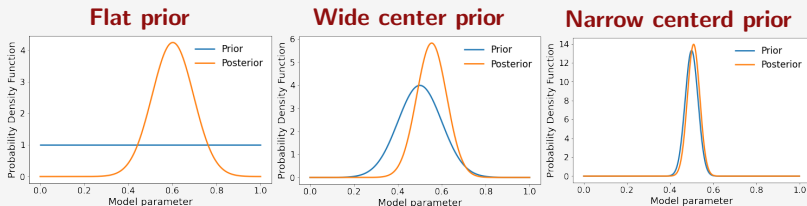
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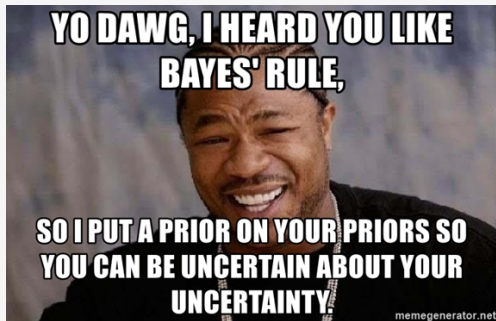
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  - expect it depends on the **choice of the prior ...**



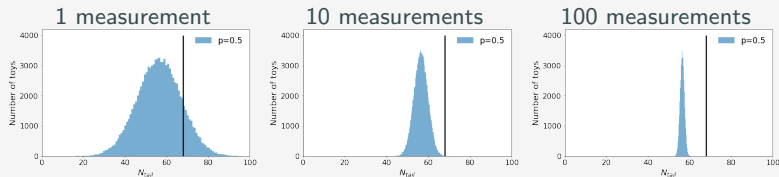
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## Frequentist



**Frequentists say “Yes, the coin is tricked!”**

Certainty comes from the extremely low fraction of pseudo-experiments of a normal coin, that would lead the observed result.

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- prior is build while accumulating knowledge, **supressing the arbitrariness**



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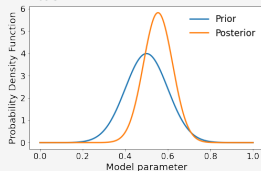
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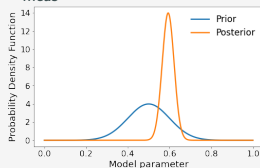
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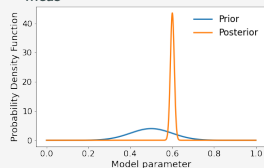
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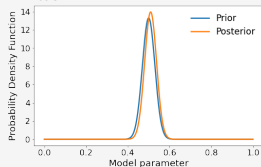
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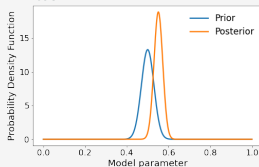
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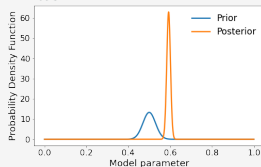
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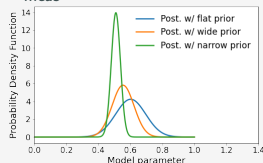
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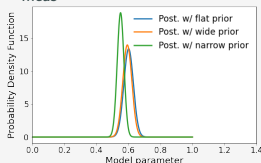
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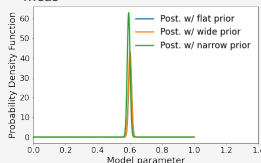
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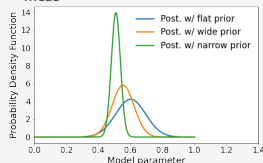
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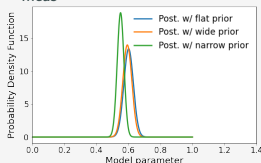
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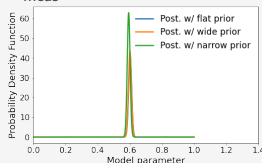
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Bayesians also say “Yes, the coin is tricked!”

## Frequentist v.s. Bayesian: what to take away

1. Both approaches handle differently the “non fully certain”
2. Final conclusions should be compatible, even if the question they address are not exactly the same.
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## One thing I like from the two approaches

- probability interpretation from the frequentist
- ranking two theories using their probability, called Bias factors

## Part I: statistics

descriptive statistics – sample – mean – (co)variance – (de)correlation

## Part II: probability

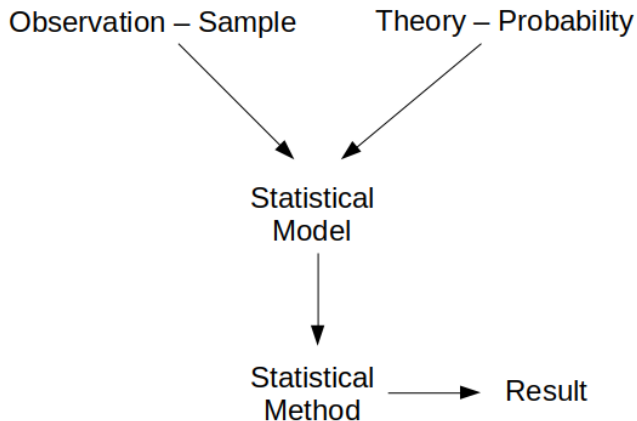
Bias theorem – prior – posterior – random variable – (marginal) PDF – moments – characteristic function – (in)dependent variables – CLT – error propagation

## Part III: statistical model

Likelihood – nuisance parameter – parameter of interest – systematic uncertainties

## Part IV: The two big school

Frequentist – occurrence frequency – pseudo-data (toys) – bayesian – degree of belief



1. **Statistics**
2. **Probability**
3. **Statistical model**
4. **The two big schools**
5. **Parameter estimation and hypothesis testing**

# Parameter estimation and hypothesis testing

---

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Baics of **parameter estimation** in both **frequentist** and **bayesian**,  
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## 3. Coming back on nuisance parameters (*i.e.* uncertainties on the model)

## Frequentist approach: estimators

**Definition:** random variable which gives a 'good' estimate of your parameter of interest ( $\hat{\mu} = \frac{1}{N} \sum_i x_i$  as estimator of  $\mathbb{E}[X]$ ). Estimator depends on observation  $\hat{\mu}(x_1, \dots, x_n)$  and is *not* constant.  $N_{meas}$  needed to assess its quality.

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## Two important examples of estimators

1. Maximum likelihood estimator (MLE):  $\hat{\mu}$  which maximizes  $\mathcal{L}(\mu; x)$   
 $\rightarrow$  numerically easier to minimize  $-2 \ln \mathcal{L}(\mu; x)$  - **negative log likelihood (NLL)**
2.  $\chi^2$  estimator:  $\hat{\mu}$  which minimizes  $\chi^2(\mu) \equiv \sum_i w_i (X_i^{pred}(\mu) - x_i)^2$

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## Question 2 for the audience:

Why consistency and bias of an estimator are different?

## Example: linear fit

**Model**  $N^{pred}(p_0, p_1; t) = p_0 + p_1 t$

**4 estimators** (or “cost function”) are used:

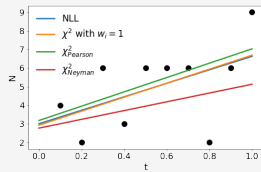
$$-2 \log \mathcal{L}_{poisson}$$

$$\chi^2(p_0, p_1) = \sum_i (N_i^{pred}(p_0, p_1) - N_i)^2$$

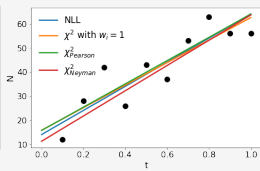
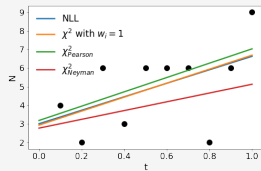
$$\chi_{Pearson}^2(p_0, p_1) = \sum_i \left( \frac{N_i^{pred}(p_0, p_1) - N_i}{\sqrt{N_i^{pred}(p_0, p_1)}} \right)^2$$

$$\chi_{Neyman}^2(p_0, p_1) = \sum_i \left( \frac{(N_i^{pred}(p_0, p_1) - N_i)^2}{\sqrt{N_i}} \right)^2$$

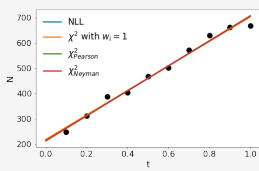
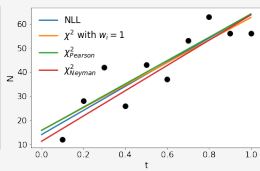
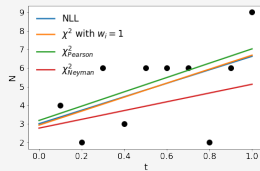
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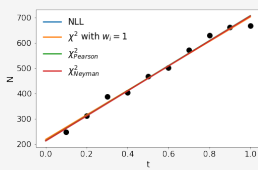
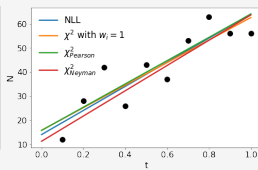
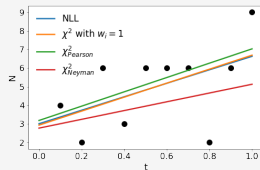
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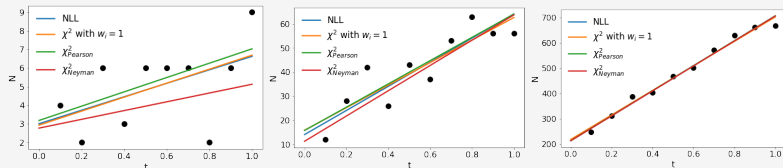


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- $\chi^2_{\text{pearson}} \equiv -2 \log \mathcal{L}_{\text{Gauss}} \approx -2 \log \mathcal{L}_{\text{Poiss}}$  for large numbers



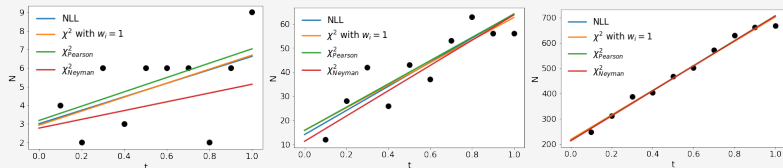
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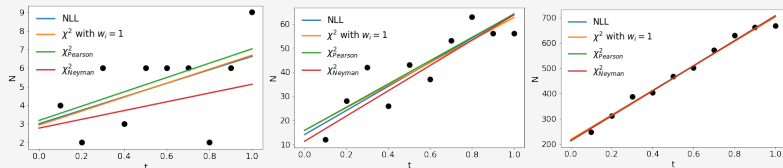
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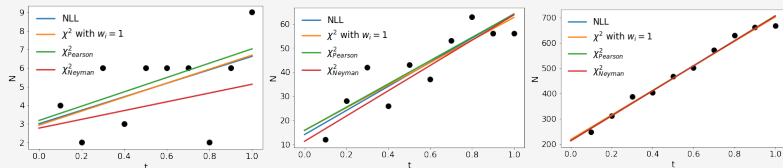
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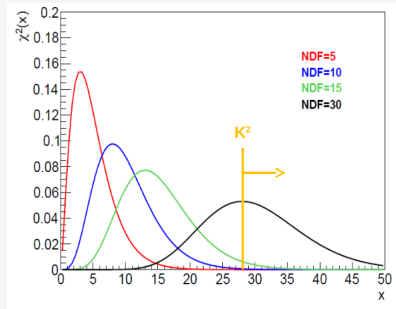
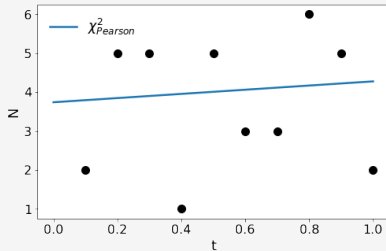
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# The basics of goodness-of-fit

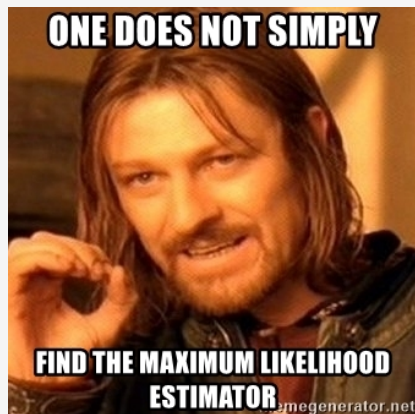


$\chi^2_{\min} = 6.7$  with 10 data points ( $nDoF = 10$ )  $\rightarrow$  blue PDF tells us this is a **good fit**, even if not a point is on the line.

We can actually compute the **fraction of pseudo-data** that would lead to a **higher  $\chi^2$**  ( $p$ -value), to **quantify** this statement.

1. Perform a fit of an histogram in ROOT, with quite wide binning. Do you recover the true value? Does the result depends on the number of bins? How to solve it?

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2. Imagine you have one dataset, but you want to fit simultaneously two distributions of these events. How to write the  $\chi^2$ ?





## Confidence interval and level $\mu \in [\mu_{min}, \mu_{max}] @ \alpha CL$

- $\equiv$  the true value is in  $[\mu_{min}, \mu_{max}]$  in  $\alpha\%$  of all possible realisations
- $\mu_{min}$  ( $\mu_{max}$ ) is the lower (upper) bound
- $\alpha$  is the confidence level
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$n$  is called “number of  $\sigma$ ” and  $\alpha(n)$  is known for a normal PDF:

- $\alpha(1) = 68\%$
- $\alpha(1.64) = 90\%$
- $\alpha(1.95) = 95\%$
- $\alpha(2) = 95.4\%$
- $\alpha(3) = 99.7\%$
- $\alpha(5) = 99.99994\%$

## Quality of a given confidence interval

- CI  $\equiv$  random variable: consider the limit of  $\infty$  number of meas.
- Coverage  $\equiv$  probability  $P$  that the true parameter *actually is* in C
- “Confidence level = what we target” while “coverage = what we get”

## The 3 cases

1.  $P = \alpha$  : perfect coverage  $\rightarrow$  ideal
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**In practice:** **estimating coverage** can be done using **toys experiment** (CPU-intensive for realistic models).

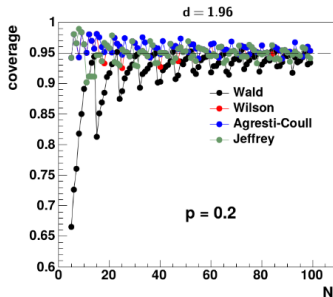
# Frequentist parameter uncertainty

**Example:** binomial distribution, with parameter of interest  $p$

$$P(k; N, p) = \binom{N}{k} p^k (1-p)^{N-k}$$

$$\hat{p} = \frac{k}{N}$$

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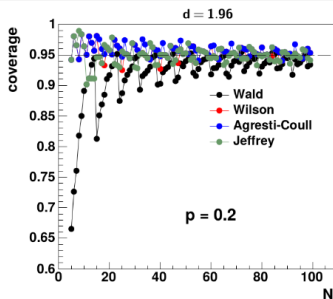
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**Take away messages:**

- notation  $\mu = X_{-Z}^{+Y}$  (assuming 68% C.L.) is sometimes only *indicative*
- *only object* which contains the *full* information is *likelihood*
- OK to manipulate these approximate quantities - *just know what they are*(n't)



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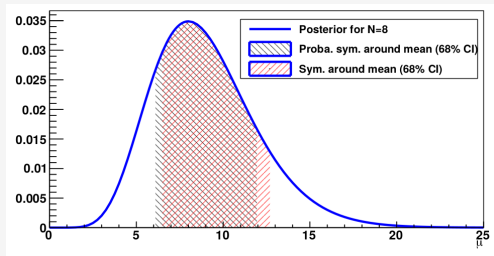
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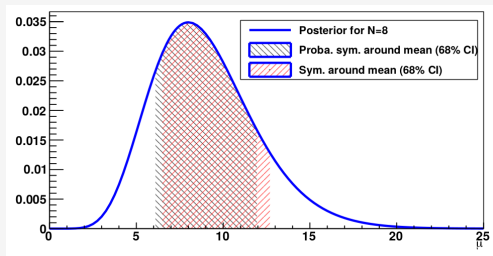
- Replace  $\mathbb{E}[\mu]$  by the mode, or the median ...

# Bayesian parameter estimation





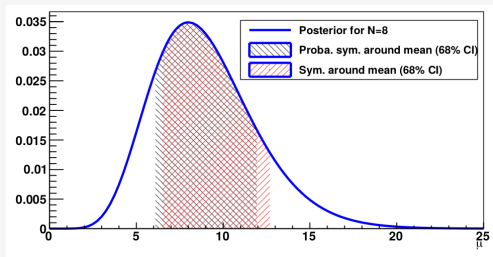
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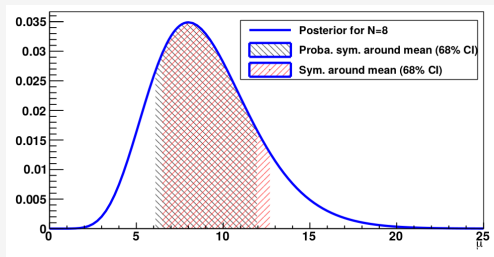
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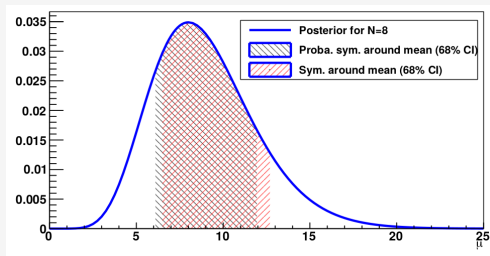
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- questions: (1) **why there is no coverage in bayesian?**  
(2) **Why the 3 properties of frequentist estimator are defined in bayesian?**

**Frequentist approach** imagine you measure energy response  $r_E$  of a detector using a dedicated data  $d_E$

- this measure is described by a likelihood  $\mathcal{L}_{energy}(r_E, d_E)$
- the parameter of interest will be better known with more data
- this unknown can be added to the stat model using the **full** likelihood

$$\mathcal{L}(\mu, r_E; data, d_E) = \mathcal{L}(\mu, ; data)\mathcal{L}_{energy}(r_E, d_E)$$

- this is notion of **auxiliary measurement**.
- $\mathcal{L}_{energy}(r_E, d_E)$  is usually too complex to be implemented.
- One uses its approximation (Taylor Expansion of order 2 of NLL around the min, leading to a gaussian likelihood)

## Coming back to model uncertainties - II

**Bayesian approach** imagine you have a calculation with some approximations, to which an uncertainty is associated.

- this uncertainty is closer to a degree of belief
- a prior  $\pi(\theta)$  is required to quantify, were the true value of  $\theta$  is more likely to be
- this unknown can be added to the stat model using the **full** likelihood

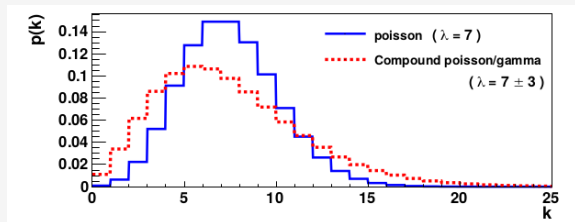
$$\mathcal{L}(\mu, \theta; data) = \mathcal{L}(\mu, ; data) \pi(\theta)$$

- this final likelihood is **marginalized** over  $\theta$ :

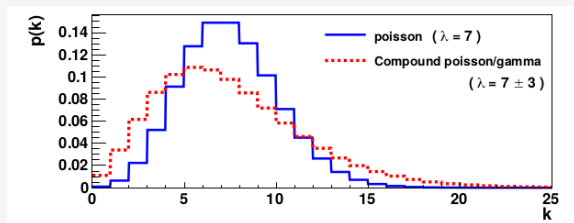
$$\mathcal{L}_m(\mu; data) = \int \mathcal{L}(\mu, \theta; data) \pi(\theta) d\theta$$

- Interpretation: average all possible situations (defined by a  $\theta$  value), accounting for the probability to actually have this value

## Example of marginalization



## Example of marginalization



## What's the proper way to implement uncertainties?

- no absolute answer to this question  $\rightarrow$  arbitrariness
- make your choice depending on the context (ease interpretation or calculation, or ...?)
- always check the robustness of your conclusion wrt these choices



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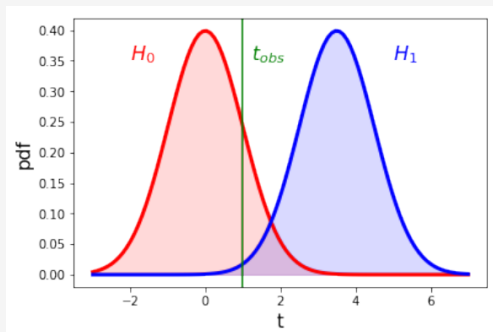
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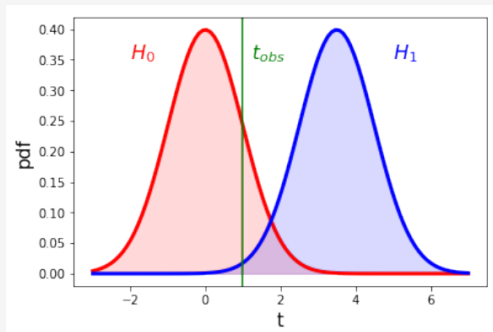
**Most naive approach:** event count as test statistics  $t = N$

- e.g.  $H_1$  predicts  $N_1 = 110$ , while  $H_0$  predicts  $N_1 = 100$
- observation  $N_{obs} = 112$ : do I reject the signal hypothesis?
- Steps of test hypothesis
  - find distribution of  $t$  in both hypothesis  $f(t|H_0)$  and  $f(t|H_1)$
  - check where  $t_{obs}$  fall wrt to  $f(t|H_0)$  and  $f(t|H_1)$
  - conclude with a confidence level ( $p$ -value)

# Test of Hypothesis



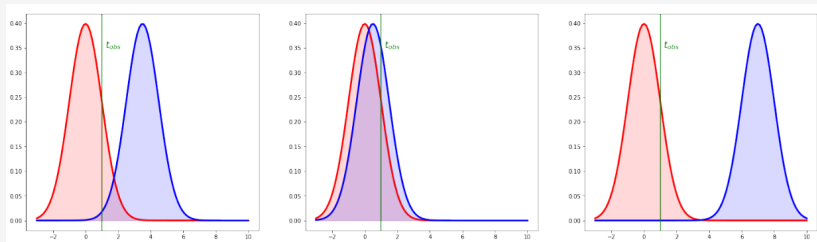
# Test of Hypothesis



**Quantitative agreement with an hypothesis:**  $p$ -value

$p$ -value = probability to observe what you observed in measurement or "more extreme" values

## How to find exclusion limit



→ Increase the signal until the signal hypothesis get rejected (at a given confidence level).



Egon Pearson



Jerzy Neyman

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- the most powerful statistical test is **Negative Log Likelihood ratio**

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In practice: **hundreds or thousands of event counts!**

## Part I: statistics

descriptive statistics – sample – mean – (co)variance – (de)correlation

## Part II: probability

Bias theorem – prior – posterior – random variable – (marginal) PDF – moments – characteristic function – (in)dependent variables – CLT – error propagation

## Part III: statistical model

Likelihood – nuisance parameter – parameter of interest – systematic uncertainties

## Part IV: The two big school

Frequentist – occurrence frequency – pseudo-data (toys) – bayesian – degree of belief

## Part VI: Parameter estimation & hypothesis testing

estimator and its properties –  $\chi^2$  – confidence/credibility level/interval – coverage –  $p$ -value – LLR

## Concluding remarks

Statistics deals with the 'not fully known'  
→ not a single way → some arbitrariness

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**Ernest Rutherford**

"If your experiment needs a statistician, you need a better experiment"

**Thanks for you attention !**