# Statistical fluctuations and artificial constraints on systematic uncertainties

## Andrey Popov On behalf of CMS collaboration

Université libre de Bruxelles

IRN Terascale Brussels, 16–18 Oct 2019

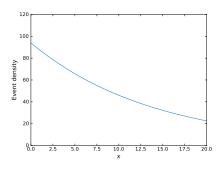






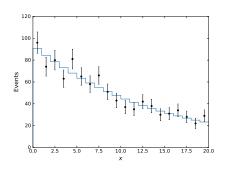


• Choose a distribution to study



- Choose a distribution to study
- Represented by a histogram

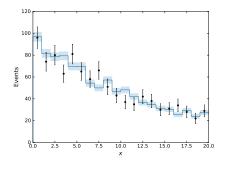
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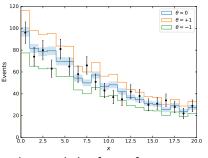
- Expected distribution ('template') constructed using Monte–Carlo
  - Associated per-bin uncertainties  $\sigma_i$



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- Expected distribution ('template') constructed using Monte–Carlo
  - Associated per-bin uncertainties  $\sigma_i$
- Systematic variations given by alternative templates
  - Nuisance parameters to control inter- and extrapolation from reference templates:  $\lambda_i = \lambda_i(\theta; \lambda_i^0, \lambda_i^+, \lambda_i^-)$



#### Likelihood

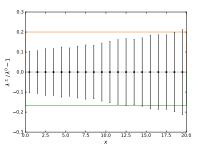
Likelihood for toy model with one physical systematic uncertainty:

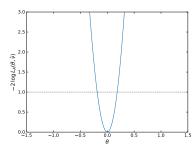
$$\log L(\theta, \boldsymbol{\nu}; \mathbf{n}) = \sum_{i=1}^{m} \log P(n_i; \lambda_i(\theta) + \nu_i \sigma_i) - \theta^2/2 - \boldsymbol{\nu}^2/2 + \text{const}$$

- Poissonian term:  $\log P(n; \lambda) = n \log \lambda \lambda \text{const}$
- $\circ$  Nuisances u control variations due to per-bin MC stat. uncertainties
- Can maximize  $\log L$  with respect to u analytically
  - Barlow–Beeston light method
- · Sensitivity is typically assessed with Asimov data set
  - Set **n** to expectation, i.e.  $n_i = \lambda_i^0$

## Constraints on systematic uncertainties

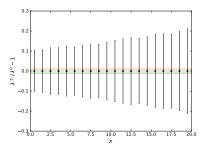
- Sensitivity to a systematic uncertainty is given by profiled likelihood
- If the variation is large compared to statistical uncertainties, it can be constrained

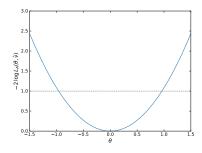




## Constraints on systematic uncertainties

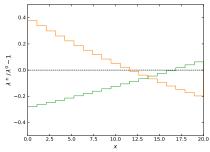
- Sensitivity to a systematic uncertainty is given by profiled likelihood
- No additional constraints if the variation is small





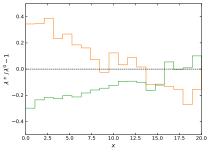
## Fluctuations in systematic variations

- There are different types of systematic variations:
  - Global or per-event weights
    - Do not change the set of MC events that enter a particular bin
    - Uncertainties in cross sections, lepton ID efficiencies, etc.



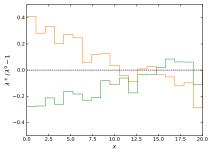
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    - Constructed from dedicated samples
    - · Some theoretical uncertainties



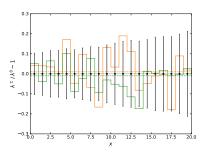
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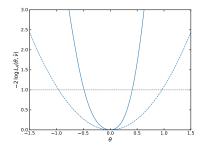
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  - Inter-bin migrations
    - Move events in and out of the signal region as well as between bins
    - · Jet momentum calibration and like



## Constraints in the presence of fluctuations

- Fluctuations in templates describing systematic uncertainties lead to tighter constraints on corresponding nuisances
  - These constraints do not represent sensitivity to underlying physical effect





## No sensitivity case

ullet As a proxy for the constraint, can use profiled likelihood at  $heta=\pm 1$ 

$$\log R \equiv \max_{\nu} \log \frac{L_A(\pm 1, \nu)}{L_A^{\text{max}}} = \sum_{i=1}^m \log \frac{P(\lambda_i^0; \lambda_i^{\pm} + \hat{\nu}_i \sigma_i)}{P(\lambda_i^0; \lambda_i^0)} - \hat{\nu}^2 / 2 - 1/2$$

- If  $\log R < -1/2$ , there is an additional constraint on  $\theta$
- $\circ$  Assuming parabolic dependence,  $| heta| < (-2 \log R)^{-1/2}$  at 68% CL
- With  $|\Delta \lambda_i| \ll \lambda_i^0$  and setting  $k = \mathcal{L}_{ ext{eff}}^{ ext{MC}}/\mathcal{L}^{ ext{Data}}$ ,

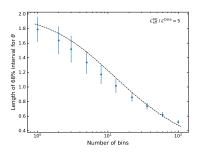
$$\log R \approx -\frac{1}{2(k+1)} \sum_{i=1}^{m} \frac{(\lambda_i^{\pm} - \lambda_i^{0})^2}{\sigma_i^2} - 1/2$$

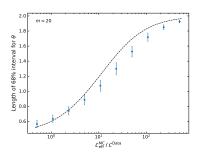
• If there is no real sensitivity, then  $\lambda_i^{\pm} \sim \mathcal{N}(\lambda_i^0, \sigma_i^2)$ , the sum follows  $\chi_m^2$  distribution, and

$$\langle \log R \rangle = -\frac{m}{2(k+1)} - 1/2$$

## No sensitivity case

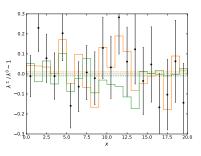
- Numeric study reproduces analytical results (shown with dashed lines)
  - The constraints become arbitrary tight as the number of bins grows
  - $\circ$  Even with  $\mathcal{O}(10)$  bins, impractically large  $\mathcal{L}_{\text{eff}}$  might be needed to avoid the constraints

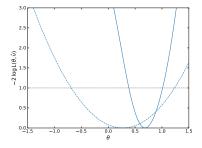




#### Constraints with data

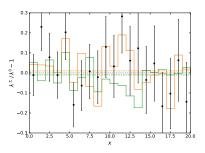
 The constraints are not an artifact of Asimov data set and occur also with pseudodata:

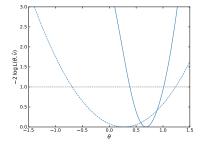




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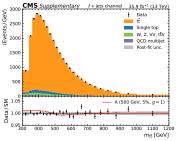


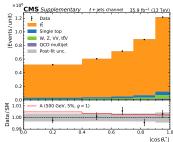


- Fit finds  $\hat{\theta}$  such that  $\lambda(\theta)$  resembles noise in data best. Any deviations from  $\hat{\theta}$  get penalized by the  $\nu^2/2$  term in the likelihood
  - Parameters u have to be adjusted for a larger difference between the noise patters in data and in  $\lambda(\theta)$

## Real-life example: $H \rightarrow t\bar{t}$

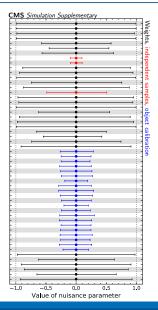
- CMS search for heavy Higgs boson decaying to  $t\bar{t}$ 
  - o arXiv:1908.01115, submitted to JHEP
- Focus on  $\ell$  + jets channel as an example
  - Reconstructed  $m_{t\bar{t}}$  as main observable
  - Angular variable reflects spin
  - Use 2D distribution in statistical analysis
    - Separate e + jets and  $\mu + jets$  channels
    - 25 × 5 bins in each channel



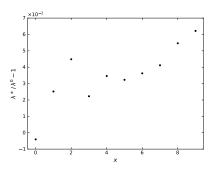


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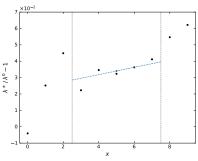
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    - Separate e + jets and  $\mu + \text{jets}$  channels
    - 25 × 5 bins in each channel
- Initially observed unexpectedly tight constraints on some of nuisances
  - Asimov data set, MC stat. uncertainties not shown
  - All these templates affected by fluctuations



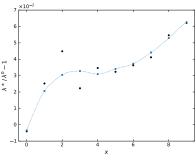
- Suppress fluctuations by smoothing relative deviations  $\lambda_i^\pm/\lambda_i^0-1$ 
  - o Denser binning than used in the analysis to avoid binning artefacts
  - $\circ$  Assume relative deviations are identical between  $e+{
    m jets}$  and  $\mu+{
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  - o Assume up and down deviations are symmetric in shape



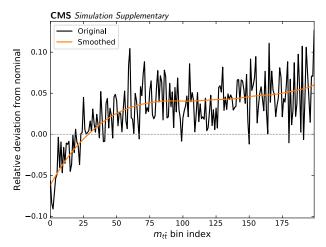
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- Local linear regression as smoothing algorithm (LOWESS)
  - Weighted least squares fit to 2D rel. deviation with a linear function
    - Restrict the fit to rectangular window
    - Points further away from the centre of window receive smaller weights



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- Local linear regression as smoothing algorithm (LOWESS)
  - Weighted least squares fit to 2D rel. deviation with a linear function
    - Restrict the fit to rectangular window
    - Points further away from the centre of window receive smaller weights
  - Repeat while sliding the window
- In each bin, apply smoothed relative deviation to the nominal template to construct new systematic variations
  - $\circ$  Rescaled independently for up and down variations minimizing  $\chi^2$  error

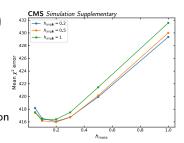


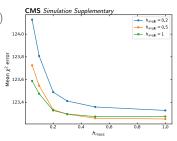
• Example of smoothing (a single angular bin shown)



## Parameters of smoothing algorithm

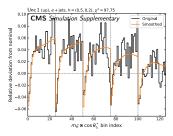
- Dimensions of the window ('bandwidths') are free parameters of the algorithm
- Chosen with repeated cross-validation
  - Split events into k = 10 partitions
  - $\circ$  Build smoothed rel. deviations from first k-1 partitions and compute approximation  $\chi^2$  error on the  $k^{\rm th}$  partition
  - Repeat for the other k-1 possible choices of the test partition
  - Repeat everything with different (random) splittings of events into partitions
  - Choose bandwidths that give smallest average approximation error

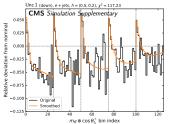


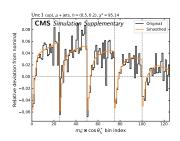


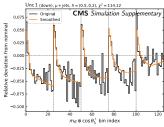
## Example of smoothing

#### An uncertainty with a real impact



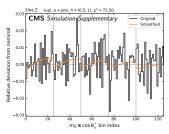


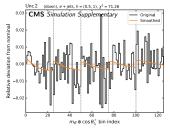


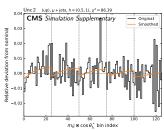


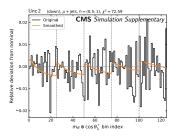
## Example of smoothing

An uncertainty with very little real impact

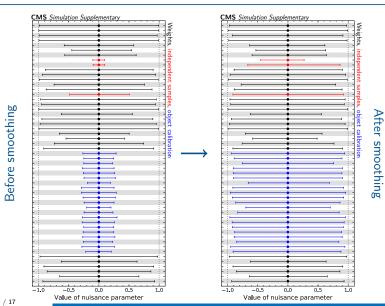








#### Constraints



## Summary

- Statistical fluctuations in templates representing systematic uncertainties can lead to severe unphysical constraints on corresponding nuisances
  - $\circ~$  Get tighter as the number of bins grows or  $\mathcal{L}_{eff}^{MC}/\mathcal{L}^{Data}$  decreases
    - Becomes especially important as large data sets are collected
  - Not addressed by MC stat. uncertainties in nominal templates
- Intuitive explanation: Fit takes into account not only the physical variation but also similarity between noise patterns in data and MC
- Smoothing of variations w. r. t. the nominal template can lift or reduce the constraints
  - Local regression was used in CMS search for  $H \to t\bar{t}$ , but other options are possible (e. g. smoothing splines or even (regularized) polynomial fit)